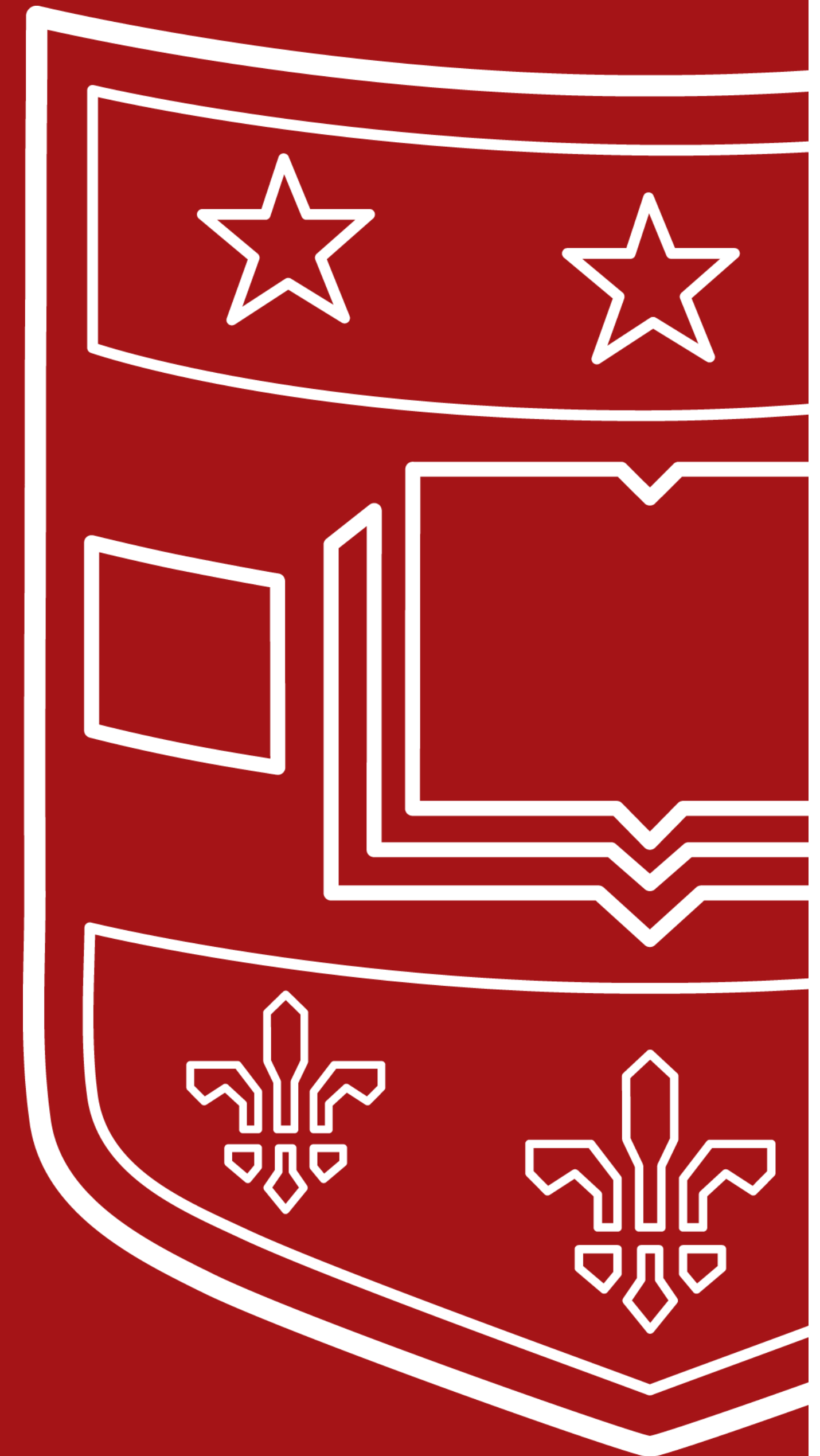


Bayes of My Life: Bayesian analysis of nucleon-nucleon scattering data in pionless EFT

Jason Bub

In collaboration with: Maria Piarulli, Dick Furnstahl,
Saori Pastore, and Daniel Phillips

The Nuclear Interaction: Post-Modern Developments @ ECT*
20 August 2024



Outline

- Motivation/Bayesian EFT Model Calibration
- BUQEYE Formalism
- Interaction Choice
- Results
- Questions

Outline

- **Motivation/Bayesian EFT Model Calibration**
- BUQEYE Formalism
- Interaction Choice
- Results
- Questions

Next-Generation χ EFT Interactions



[arXiv:2408.02480](https://arxiv.org/abs/2408.02480)



Next-Generation χ EFT Interactions



We are interested in calibrating the next generation of EFT nucleon-nucleon interactions.

[arXiv:2408.02480](https://arxiv.org/abs/2408.02480)



Next-Generation χ EFT Interactions



We are interested in calibrating the next generation of EFT nucleon-nucleon interactions.

These models should have robust uncertainty quantification:

[arXiv:2408.02480](https://arxiv.org/abs/2408.02480)



Next-Generation χ EFT Interactions



We are interested in calibrating the next generation of EFT nucleon-nucleon interactions.

These models should have robust uncertainty quantification:

- Parametric uncertainty

[arXiv:2408.02480](https://arxiv.org/abs/2408.02480)



Next-Generation χ EFT Interactions



We are interested in calibrating the next generation of EFT nucleon-nucleon interactions.

These models should have robust uncertainty quantification:

- Parametric uncertainty
- Truncation uncertainty

[arXiv:2408.02480](https://arxiv.org/abs/2408.02480)



Next-Generation χ EFT Interactions



We are interested in calibrating the next generation of EFT nucleon-nucleon interactions.

These models should have robust uncertainty quantification:

- Parametric uncertainty
- Truncation uncertainty

[arXiv:2408.02480](https://arxiv.org/abs/2408.02480)

This must be accomplished in the model calibration.



Model Calibration



Model Calibration



We have LECs that are unknown.

Model Calibration



We have LECs that are unknown.

How do we find them?

Model Calibration



We have LECs that are unknown.

How do we find them?

- Compute on the lattice?

Model Calibration



We have LECs that are unknown.

How do we find them?

- Compute on the lattice? ← Too hard



Model Calibration

We have LECs that are unknown.

How do we find them?

- Compute on the lattice? ← Too hard

So how do we **calibrate** the LECs?



Model Calibration

We have LECs that are unknown.

How do we find them?

- Compute on the lattice? ← Too hard

So how do we **calibrate** the LECs?

Let's cast the problem this way:

$$\text{pr}(\vec{a} | \vec{y}, I) \propto \text{pr}(\vec{y} | \vec{a})$$



Model Calibration

We have LECs that are unknown.

How do we find them?

- Compute on the lattice? ← Too hard

So how do we **calibrate** the LECs?

Let's cast the problem this way:

$$\text{pr}(\vec{a} | \vec{y}, I) \propto \text{pr}(\vec{y} | \vec{a})$$

where \vec{a} is the vector of parameters, \vec{y} is a vector of data, and I is additional external information.

Likelihood Modeling



Likelihood Modeling



We have

Likelihood Modeling



We have

$$\underbrace{\text{pr}(\vec{a} \mid \vec{y}, I)}_{\text{Posterior}} \propto \underbrace{\text{pr}(\vec{y} \mid \vec{a})}_{\text{Likelihood}}$$

Likelihood Modeling



We have

$$\underbrace{\text{pr}(\vec{a} \mid \vec{y}, I)}_{\text{Posterior}} \propto \underbrace{\text{pr}(\vec{y} \mid \vec{a})}_{\text{Likelihood}}$$

How do we calculate the likelihood?

Likelihood Modeling



We have

$$\underbrace{\text{pr}(\vec{a} \mid \vec{y}, I)}_{\text{Posterior}} \propto \underbrace{\text{pr}(\vec{y} \mid \vec{a})}_{\text{Likelihood}}$$

How do we calculate the likelihood?

Let's say we have some experimental data, \vec{y}_{exp} . For each piece of data, we assume this is given by



Likelihood Modeling

We have

$$\underbrace{\text{pr}(\vec{a} \mid \vec{y}, I)}_{\text{Posterior}} \propto \underbrace{\text{pr}(\vec{y} \mid \vec{a})}_{\text{Likelihood}}$$

How do we calculate the likelihood?

Let's say we have some experimental data, \vec{y}_{exp} . For each piece of data, we assume this is given by

$$y_{\text{exp}} = y_{\text{th}}(\vec{a}) + \delta y_{\text{exp}}, \quad \delta y_{\text{exp}} \sim \mathcal{N}(0, \sigma^2)$$

χ^2 from the Likelihood



χ^2 from the Likelihood



Since we made a Gaussian assumption for the experimental discrepancy, we can easily arrive at the result

χ^2 from the Likelihood



Since we made a Gaussian assumption for the experimental discrepancy, we can easily arrive at the result

$$\text{pr}(\mathbf{y} | \vec{\mathbf{a}}) \sim e^{-\left(y_{\text{exp}} - y_{\text{th}}(\vec{\mathbf{a}})\right)^2 / 2\sigma^2}$$

χ^2 from the Likelihood



Since we made a Gaussian assumption for the experimental discrepancy, we can easily arrive at the result

$$\text{pr}(\mathbf{y} | \vec{\mathbf{a}}) \sim e^{-\left(y_{\text{exp}} - y_{\text{th}}(\vec{\mathbf{a}})\right)^2 / 2\sigma^2}$$

For independent data, this becomes

χ^2 from the Likelihood



Since we made a Gaussian assumption for the experimental discrepancy, we can easily arrive at the result

$$\text{pr}(\mathbf{y} | \vec{\mathbf{a}}) \sim e^{-\left(y_{\text{exp}} - y_{\text{th}}(\vec{\mathbf{a}})\right)^2 / 2\sigma^2}$$

For independent data, this becomes

$$\text{pr}(\vec{\mathbf{y}} | \vec{\mathbf{a}}) \sim e^{-\sum_i \left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{\mathbf{a}})\right)^2 / 2\sigma_i^2}$$

χ^2 from the Likelihood



Since we made a Gaussian assumption for the experimental discrepancy, we can easily arrive at the result

$$\text{pr}(\mathbf{y} | \vec{\mathbf{a}}) \sim e^{-\left(y_{\text{exp}} - y_{\text{th}}(\vec{\mathbf{a}})\right)^2 / 2\sigma^2}$$

For independent data, this becomes

$$\text{pr}(\vec{\mathbf{y}} | \vec{\mathbf{a}}) \sim e^{-\sum_i \left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{\mathbf{a}})\right)^2 / 2\sigma_i^2} = e^{-\chi^2/2}$$

χ^2 from the Likelihood



Since we made a Gaussian assumption for the experimental discrepancy, we can easily arrive at the result

$$\text{pr}(\mathbf{y} | \vec{\mathbf{a}}) \sim e^{-\left(y_{\text{exp}} - y_{\text{th}}(\vec{\mathbf{a}})\right)^2 / 2\sigma^2}$$

For independent data, this becomes

$$\text{pr}(\vec{\mathbf{y}} | \vec{\mathbf{a}}) \sim e^{-\sum_i \left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{\mathbf{a}})\right)^2 / 2\sigma_i^2} = e^{-\chi^2/2}$$

We have arrived that the familiar χ^2 :

χ^2 from the Likelihood



Since we made a Gaussian assumption for the experimental discrepancy, we can easily arrive at the result

$$\text{pr}(\mathbf{y} | \vec{\mathbf{a}}) \sim e^{-\left(y_{\text{exp}} - y_{\text{th}}(\vec{\mathbf{a}})\right)^2 / 2\sigma^2}$$

For independent data, this becomes

$$\text{pr}(\vec{\mathbf{y}} | \vec{\mathbf{a}}) \sim e^{-\sum_i \left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{\mathbf{a}})\right)^2 / 2\sigma_i^2} = e^{-\chi^2/2}$$

We have arrived that the familiar χ^2 :

$$\chi^2 = \sum_i \frac{\left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{\mathbf{a}})\right)^2}{\sigma_i^2}$$

χ^2 from the Likelihood



Since we made a Gaussian assumption for the experimental discrepancy, we can easily arrive at the result

$$\text{pr}(y | \vec{a}) \sim e^{-\left(y_{\text{exp}} - y_{\text{th}}(\vec{a})\right)^2 / 2\sigma^2}$$

For independent data, this becomes

$$\text{pr}(\vec{y} | \vec{a}) \sim e^{-\sum_i \left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{a})\right)^2 / 2\sigma_i^2} = e^{-\chi^2/2}$$

We have arrived that the familiar χ^2 :

$$\chi^2 = \sum_i \frac{\left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{a})\right)^2}{\sigma_i^2}$$

χ^2 from the Likelihood



Since we made a Gaussian assumption for the experimental discrepancy, we can easily arrive at the result

$$\text{pr}(y | \vec{a}) \sim e^{-\left(y_{\text{exp}} - y_{\text{th}}(\vec{a})\right)^2 / 2\sigma^2}$$

For independent data, this becomes

Maximize $\longrightarrow \text{pr}(\vec{y} | \vec{a}) \sim e^{-\sum_i \left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{a})\right)^2 / 2\sigma_i^2} = e^{-\chi^2/2}$

We have arrived that the familiar χ^2 :

$$\chi^2 = \sum_i \frac{\left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{a})\right)^2}{\sigma_i^2}$$

χ^2 from the Likelihood



Since we made a Gaussian assumption for the experimental discrepancy, we can easily arrive at the result

$$\text{pr}(y | \vec{a}) \sim e^{-\left(y_{\text{exp}} - y_{\text{th}}(\vec{a})\right)^2 / 2\sigma^2}$$

For independent data, this becomes

Maximize \rightarrow $\text{pr}(\vec{y} | \vec{a}) \sim e^{-\sum_i \left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{a})\right)^2 / 2\sigma_i^2} = e^{-\chi^2 / 2}$ **Minimize**

We have arrived that the familiar χ^2 :

$$\chi^2 = \sum_i \frac{\left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{a})\right)^2}{\sigma_i^2}$$

χ^2 from the Likelihood



Since we made a Gaussian assumption for the experimental discrepancy, we can easily arrive at the result

$$\text{pr}(y | \vec{a}) \sim e^{-\left(y_{\text{exp}} - y_{\text{th}}(\vec{a})\right)^2 / 2\sigma^2}$$

For independent data, this becomes

Maximize $\rightarrow \text{pr}(\vec{y} | \vec{a}) \sim e^{-\sum_i \left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{a})\right)^2 / 2\sigma_i^2} = e^{-\chi^2/2}$ **Minimize**
 \rightarrow Find best model

We have arrived that the familiar χ^2 :

$$\chi^2 = \sum_i \frac{\left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{a})\right)^2}{\sigma_i^2}$$

Bayes' Theorem



Bayes' Theorem



But this isn't the full posterior. Bayes' Theorem yields

Bayes' Theorem



But this isn't the full posterior. Bayes' Theorem yields

$$\underbrace{\text{pr}(\vec{a} | \vec{y}, I)}_{\text{Posterior}} \propto \underbrace{\text{pr}(\vec{y} | \vec{a})}_{\text{Likelihood}} \underbrace{\text{pr}(\vec{a} | I)}_{\text{Prior}}$$



Bayes' Theorem

But this isn't the full posterior. Bayes' Theorem yields

$$\underbrace{\text{pr}(\vec{a} | \vec{y}, \mathbf{I})}_{\text{Posterior}} \propto \underbrace{\text{pr}(\vec{y} | \vec{a})}_{\text{Likelihood}} \underbrace{\text{pr}(\vec{a} | \mathbf{I})}_{\text{Prior}}$$

What we found for a **Least-squares optimization** essentially had $\text{pr}(\vec{a} | \mathbf{I}) = \mathbf{C}$.



Bayes' Theorem

But this isn't the full posterior. Bayes' Theorem yields

$$\underbrace{\text{pr}(\vec{a} | \vec{y}, I)}_{\text{Posterior}} \propto \underbrace{\text{pr}(\vec{y} | \vec{a})}_{\text{Likelihood}} \underbrace{\text{pr}(\vec{a} | I)}_{\text{Prior}}$$

What we found for a **Least-squares optimization** essentially had $\text{pr}(\vec{a} | I) = C$.

What the **prior** does for us is **encode any previous information** that we may know.



Bayes' Theorem

But this isn't the full posterior. Bayes' Theorem yields

$$\underbrace{\text{pr}(\vec{a} | \vec{y}, \mathbf{I})}_{\text{Posterior}} \propto \underbrace{\text{pr}(\vec{y} | \vec{a})}_{\text{Likelihood}} \underbrace{\text{pr}(\vec{a} | \mathbf{I})}_{\text{Prior}}$$

What we found for a **Least-squares optimization** essentially had $\text{pr}(\vec{a} | \mathbf{I}) = \mathbf{C}$.

What the **prior** does for us is **encode any previous information** that we may know.

- Ex: LECs are natural, i.e., order 1 $\rightarrow \text{pr}(\vec{a} | \mathbf{I}) \sim \mathcal{N}(\vec{0}, \Sigma_{\text{pr}})$

Parameter Estimation vs. Parameter Fitting



Parameter Estimation vs. Parameter Fitting



Bayesian calibration with $\text{pr}(\vec{a} | \vec{y}, \mathbf{I}) \propto \text{pr}(\vec{y} | \vec{a}) \text{pr}(\vec{a} | \mathbf{I})$ is vastly different from **least-squares**



Parameter Estimation vs. Parameter Fitting

Bayesian calibration with $\text{pr}(\vec{a} | \vec{y}, \mathbf{I}) \propto \text{pr}(\vec{y} | \vec{a}) \text{pr}(\vec{a} | \mathbf{I})$ is vastly different from **least-squares**

Bayesian calibration

Least-squares



Parameter Estimation vs. Parameter Fitting

Bayesian calibration with $\text{pr}(\vec{a} | \vec{y}, I) \propto \text{pr}(\vec{y} | \vec{a}) \text{pr}(\vec{a} | I)$ is vastly different from **least-squares**

Bayesian calibration

- Parameters are random variables
- Parametric uncertainty

Least-squares

- Parameters are fixed



Parameter Estimation vs. Parameter Fitting

Bayesian calibration with $\text{pr}(\vec{a} | \vec{y}, I) \propto \text{pr}(\vec{y} | \vec{a}) \text{pr}(\vec{a} | I)$ is vastly different from **least-squares**

Bayesian calibration

- Parameters are random variables
- Parametric uncertainty
- Use of prior information

Least-squares

- Parameters are fixed
- Informed only by data



Parameter Estimation vs. Parameter Fitting

Bayesian calibration with $\text{pr}(\vec{a} | \vec{y}, \mathbf{I}) \propto \text{pr}(\vec{y} | \vec{a}) \text{pr}(\vec{a} | \mathbf{I})$ is vastly different from **least-squares**

Bayesian calibration

- Parameters are random variables
- Parametric uncertainty
- Use of prior information
- Robust handling of data outliers

Least-squares

- Parameters are fixed
- Informed only by data
- Sensitive to data outliers



Parameter Estimation vs. Parameter Fitting

Bayesian calibration with $\text{pr}(\vec{a} | \vec{y}, I) \propto \text{pr}(\vec{y} | \vec{a}) \text{pr}(\vec{a} | I)$ is vastly different from **least-squares**

Bayesian calibration

- Parameters are random variables
- Parametric uncertainty
- Use of prior information
- Robust handling of data outliers
- Rigorous uncertainties

Least-squares

- Parameters are fixed
- Informed only by data
- Sensitive to data outliers
- Approximate uncertainties



Parameter Estimation vs. Parameter Fitting

Bayesian calibration with $\text{pr}(\vec{a} | \vec{y}, I) \propto \text{pr}(\vec{y} | \vec{a}) \text{pr}(\vec{a} | I)$ is vastly different from **least-squares**

Bayesian calibration

- Parameters are random variables
- Parametric uncertainty
- Use of prior information
- Robust handling of data outliers
- Rigorous uncertainties
- Easy extension to full model uncertainties

Least-squares

- Parameters are fixed
- Informed only by data
- Sensitive to data outliers
- Approximate uncertainties
- Convoluted extension to model uncertainties

Likelihood Improvement



Likelihood Improvement



In the **likelihood**, we had the χ^2 , $\left(e^{-\chi^2/2} \right)$, but we can improve this.

Likelihood Improvement



In the **likelihood**, we had the χ^2 , $\left(e^{-\chi^2/2}\right)$, but we can improve this.

We can inform the model calibration with information about the model *itself*.

Likelihood Improvement



In the **likelihood**, we had the χ^2 , $\left(e^{-\chi^2/2}\right)$, but we can improve this.

We can inform the model calibration with information about the model *itself*.

In what way?

Likelihood Improvement



In the **likelihood**, we had the χ^2 , $\left(e^{-\chi^2/2}\right)$, but we can improve this.

We can inform the model calibration with information about the model *itself*.

In what way?

$$\chi^2 = \sum_i \frac{\left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{a})\right)^2}{\sigma_i^2}$$



Likelihood Improvement

In the **likelihood**, we had the χ^2 , $\left(e^{-\chi^2/2}\right)$, but we can improve this.

We can inform the model calibration with information about the model *itself*.

In what way?

$$\chi^2 = \sum_i \frac{\left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{a})\right)^2}{\sigma_i^2}$$



Likelihood Improvement

In the **likelihood**, we had the χ^2 , $\left(e^{-\chi^2/2}\right)$, but we can improve this.

We can inform the model calibration with information about the model *itself*.

In what way?

$$\chi^2 = \sum_i \frac{\left(y_{\text{exp}}^{(i)} - y_{\text{th}}^{(i)}(\vec{a})\right)^2}{\sigma_i^2} \longrightarrow \sigma_i^2 \rightarrow \sigma_i^2 + \sigma_{\text{th},i}^2$$

Theory Uncertainty in Calibration



Theory Uncertainty in Calibration

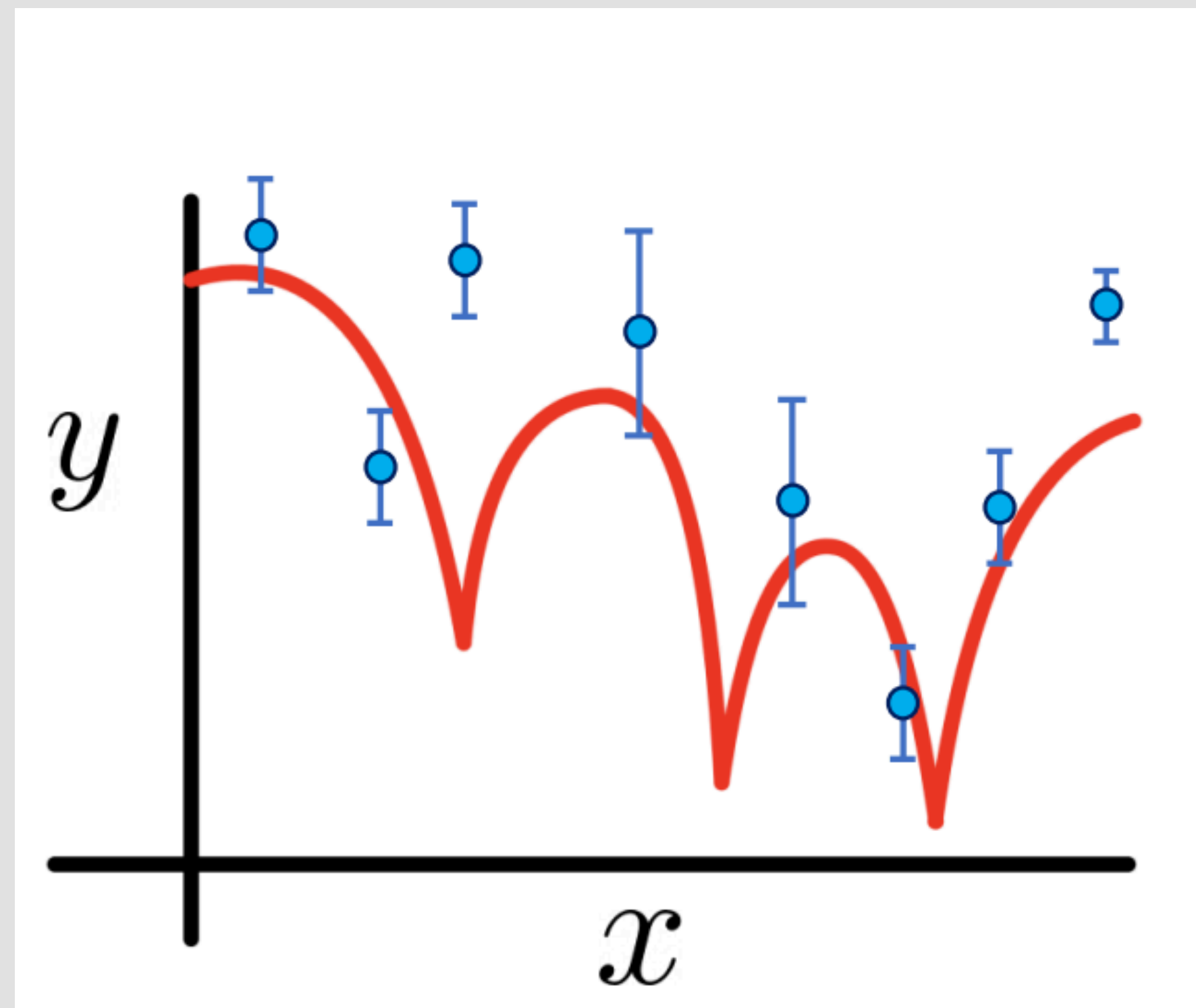


Why are theory errors necessary in calibration?

Theory Uncertainty in Calibration



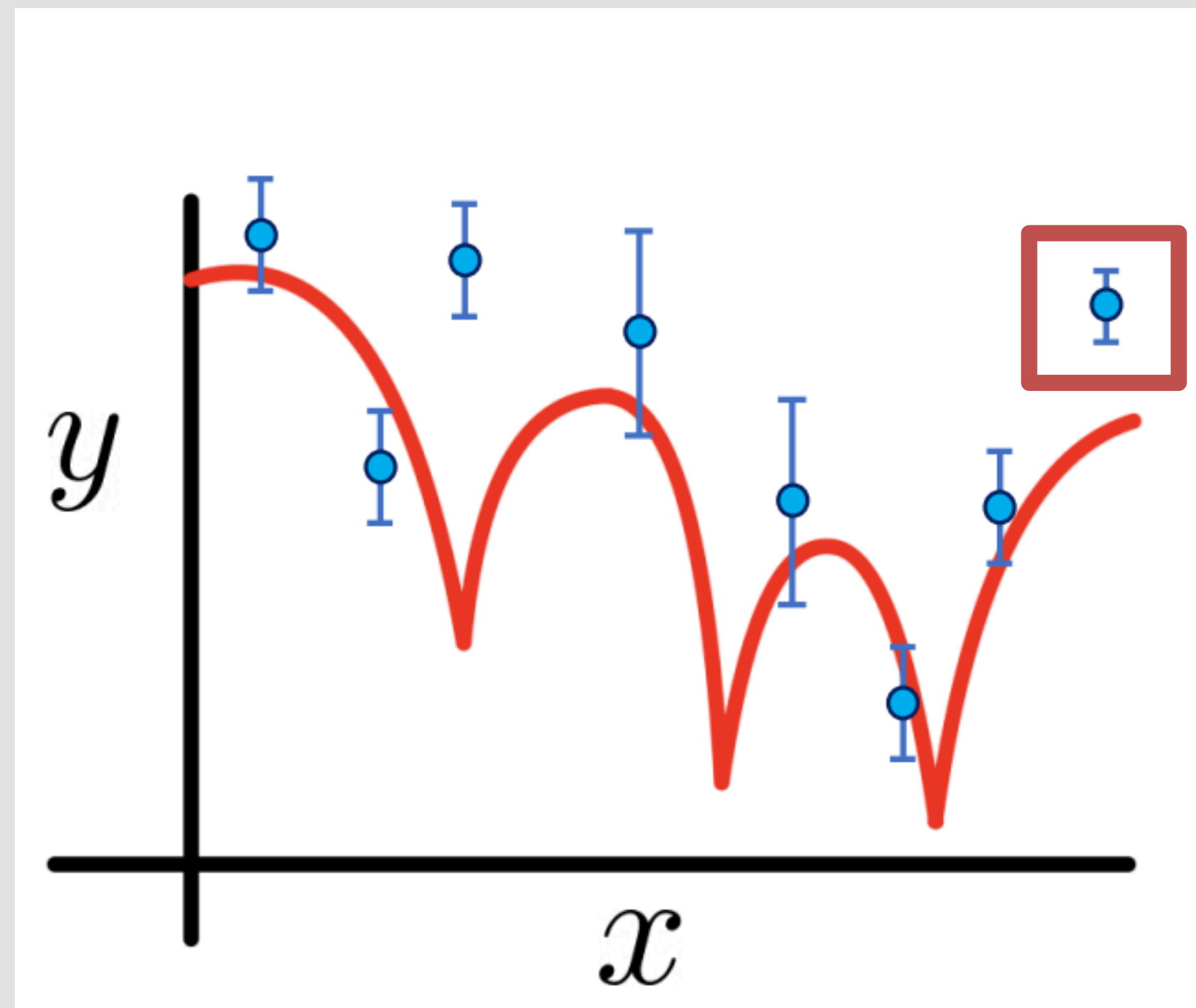
Why are theory errors necessary in calibration?



Theory Uncertainty in Calibration



Why are theory errors necessary in calibration?



Theory Uncertainty in Calibration



Why are theory errors necessary in calibration?

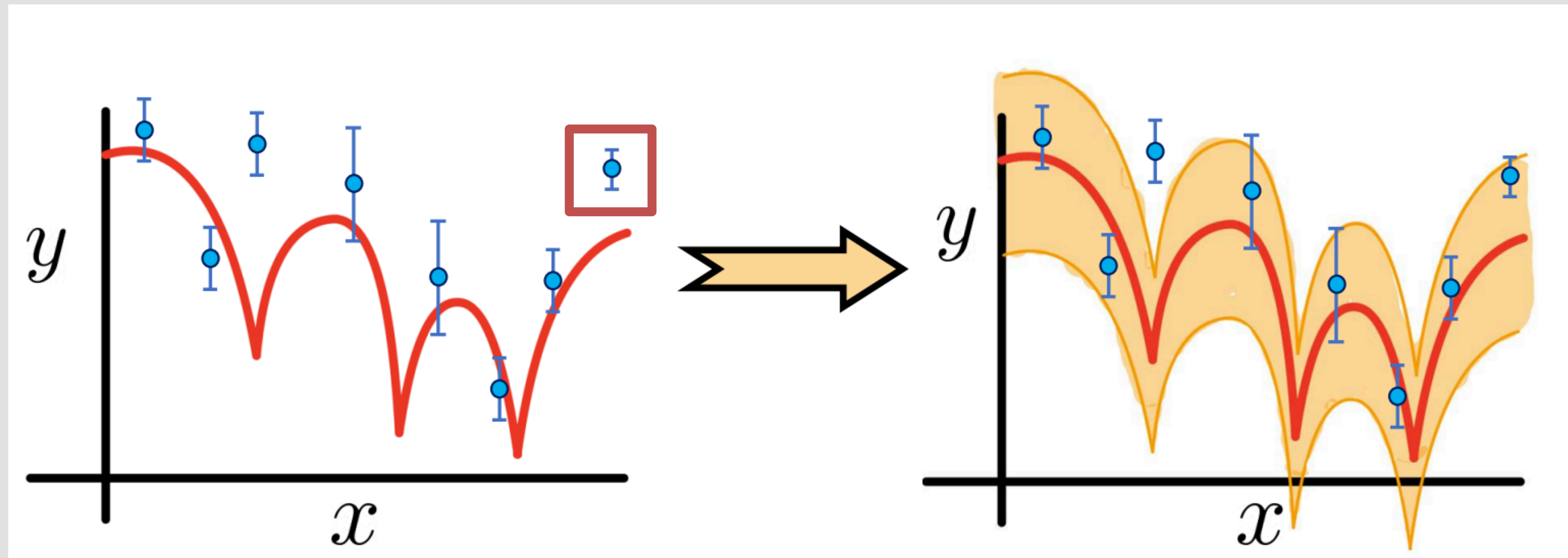


Figure courtesy of Pablo Giuliani



Theory Uncertainty in Calibration

Why are theory errors necessary in calibration?

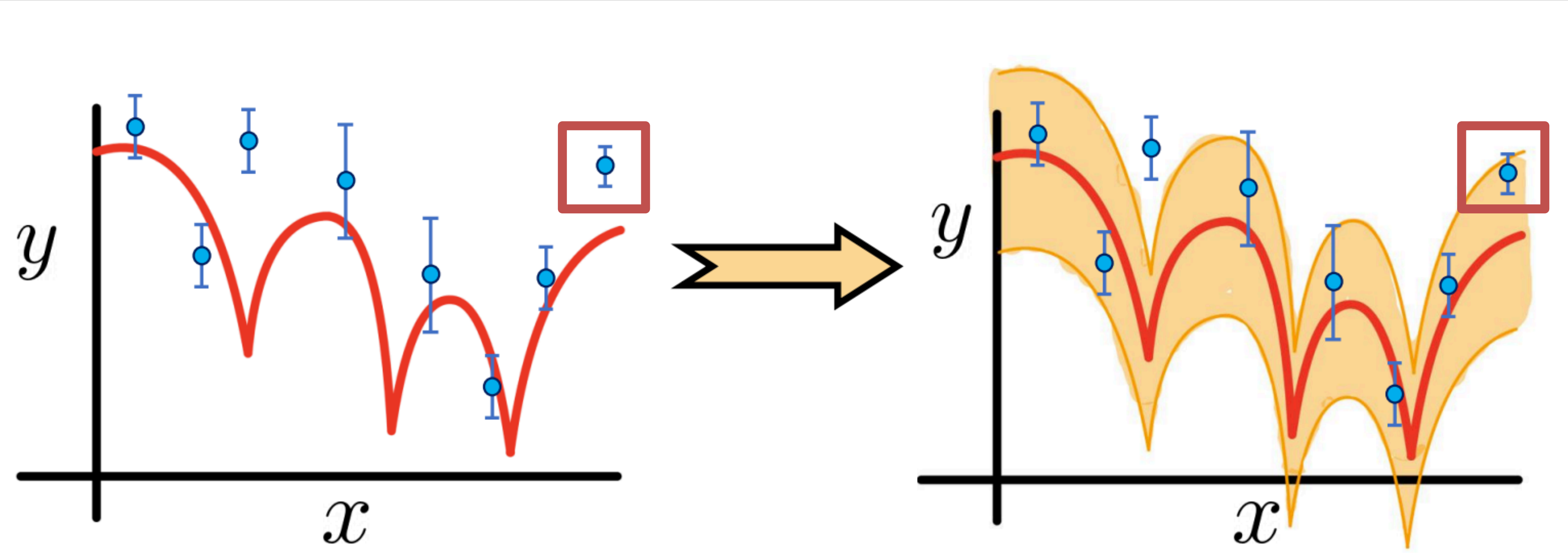


Figure courtesy of Pablo Giuliani



Theory Uncertainty in Calibration

Why are theory errors necessary in calibration?

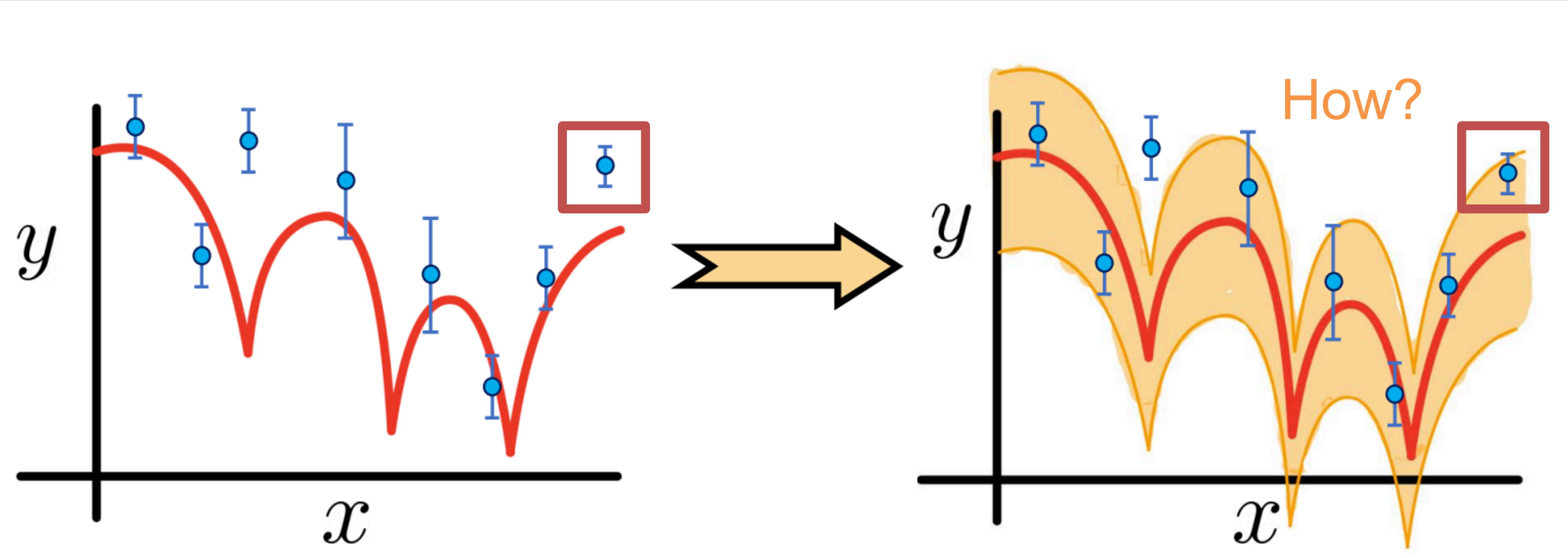


Figure courtesy of Pablo Giuliani

Outline

- Motivation/Bayesian EFT Model Calibration
- **BUQEYE Formalism**
- Interaction Choice
- Results
- Questions

Modeling the Model



Modeling the Model



Since our model is a perturbative series, we can model it as such*:

$$y_{\text{th}}(x) = y_{\text{ref}}(x) \sum_{n=0}^{\infty} c_n(x) Q^n(x), \quad Q \equiv \frac{\max[p_{\text{soft}}, p]}{\Lambda_b},$$

*R. J. Furnstahl et. al. Phys. Rev. C **92**, 024005



Modeling the Model

Since our model is a perturbative series, we can model it as such*:

$$y_{\text{th}}(x) = y_{\text{ref}}(x) \sum_{n=0}^{\infty} c_n(x) Q^n(x), \quad Q \equiv \frac{\max[p_{\text{soft}}, p]}{\Lambda_b},$$

where $y_{\text{ref}}(x)$ sets a reference scale for the observable y_{th} and Λ_b is the EFT breakdown scale.



Modeling the Model

Since our model is a perturbative series, we can model it as such*:

$$y_{\text{th}}(x) = y_{\text{ref}}(x) \sum_{n=0}^{\infty} c_n(x) Q^n(x), \quad Q \equiv \frac{\max[p_{\text{soft}}, p]}{\Lambda_b},$$

where $y_{\text{ref}}(x)$ sets a reference scale for the observable y_{th} and Λ_b is the EFT breakdown scale.

This series follows the truncation scheme of the EFT:

$$y_{\text{th}}(x) = y_{\text{ref}}(x) \sum_{n=0}^k c_n Q^n + y_{\text{ref}}(x) \sum_{n=k+1}^{\infty} c_n Q^n = y_{\text{th}}^{(k)}(x) + \delta y_{\text{th}}^{(k)}(x).$$

*R. J. Furnstahl et. al. Phys. Rev. C **92**, 024005

Truncation Errors



Truncation Errors



From the neglected terms, we have

$$\delta y_{\text{th}}^{(k)}(x) = y_{\text{ref}}(x) \sum_{n=k+1}^{\infty} c_n(x) Q^n(x).$$



Truncation Errors

From the neglected terms, we have

$$\delta y_{\text{th}}^{(k)}(x) = y_{\text{ref}}(x) \sum_{n=k+1}^{\infty} c_n(x) Q^n(x).$$

This is a geometric series in Q , so we can find*

$$\delta y_{\text{th}}^{(k)}(x) = \frac{y_{\text{ref}} \bar{c} Q^{(k+1)}}{1 - Q},$$

*J. A. Melendez et. al. Phys. Rev. C **100**, 044001



Truncation Errors

From the neglected terms, we have

$$\delta y_{\text{th}}^{(k)}(x) = y_{\text{ref}}(x) \sum_{n=k+1}^{\infty} c_n(x) Q^n(x).$$

This is a geometric series in Q , so we can find*

$$\delta y_{\text{th}}^{(k)}(x) = \frac{y_{\text{ref}} \bar{c} Q^{(k+1)}}{1 - Q},$$

Where we assume that $c_n | \bar{c} \sim \mathcal{N}(0, \bar{c}^2)$.

Theoretical Covariance





Theoretical Covariance

From the truncation uncertainty, we can construct a covariance matrix, assuming δy_{th} is normally distributed,

$$\Sigma_{ij}^{\text{th}} = \frac{\left(y_{\text{ref},i} \bar{c} Q_i^{(k+1)} \right) \left(y_{\text{ref},j} \bar{c} Q_j^{(k+1)} \right)}{1 - Q_i Q_j} r(x_i, x_j; \vec{l}),$$



Theoretical Covariance

From the truncation uncertainty, we can construct a covariance matrix, assuming δy_{th} is normally distributed,

$$\Sigma_{ij}^{\text{th}} = \frac{\left(y_{\text{ref},i} \bar{c} Q_i^{(k+1)} \right) \left(y_{\text{ref},j} \bar{c} Q_j^{(k+1)} \right)}{1 - Q_i Q_j} r(x_i, x_j; \vec{l}),$$

where we introduce a kernel $r(x_i, x_j; \vec{l})$ to smooth and handle correlations.

Correlated Likelihood



Correlated Likelihood



We can build a total covariance,

$$\Sigma_{ij} = \Sigma_{ij}^{\text{exp}} \delta_{ij} + \Sigma_{ij}^{\text{th}}.$$



Correlated Likelihood

We can build a total covariance,

$$\Sigma_{ij} = \Sigma_{ij}^{\text{exp}} \delta_{ij} + \Sigma_{ij}^{\text{th}}.$$

And our correlated likelihood is now

$$\text{pr}(\vec{y} \mid \vec{a}, \mathbf{I}) \propto e^{-\left(\vec{y}_{\text{exp}} - \vec{y}_{\text{th}}\right)^{\text{T}} \Sigma^{-1} \left(\vec{y}_{\text{exp}} - \vec{y}_{\text{th}}\right)} = e^{-d_{\text{M}}(\vec{a})}$$



Correlated Likelihood

We can build a total covariance,

$$\Sigma_{ij} = \Sigma_{ij}^{\text{exp}} \delta_{ij} + \Sigma_{ij}^{\text{th}}.$$

And our correlated likelihood is now

$$\text{pr}(\vec{y} \mid \vec{a}, \mathbf{I}) \propto e^{-\left(\vec{y}_{\text{exp}} - \vec{y}_{\text{th}}\right)^{\text{T}} \Sigma^{-1} \left(\vec{y}_{\text{exp}} - \vec{y}_{\text{th}}\right)} = e^{-d_M(\vec{a})}$$

where we define the **Mahalanobis distance**

$$d_M(\vec{a}) = \left(\vec{y}_{\text{exp}} - \vec{y}_{\text{th}}\right)^{\text{T}} \Sigma^{-1} \left(\vec{y}_{\text{exp}} - \vec{y}_{\text{th}}\right).$$



Correlated Likelihood

We can build a total covariance,

$$\Sigma_{ij} = \Sigma_{ij}^{\text{exp}} \delta_{ij} + \Sigma_{ij}^{\text{th}}.$$

And our correlated likelihood is now

$$\text{pr}(\vec{y} \mid \vec{a}, \mathbf{I}) \propto e^{-\left(\vec{y}_{\text{exp}} - \vec{y}_{\text{th}}\right)^{\text{T}} \Sigma^{-1} \left(\vec{y}_{\text{exp}} - \vec{y}_{\text{th}}\right)} = e^{-d_M(\vec{a})}$$

where we define the **Mahalanobis distance**

$$d_M(\vec{a}) = \left(\vec{y}_{\text{exp}} - \vec{y}_{\text{th}}\right)^{\text{T}} \Sigma^{-1} \left(\vec{y}_{\text{exp}} - \vec{y}_{\text{th}}\right).$$

Correlated version of χ^2

Additional Parameters



Additional Parameters



In this process, we have introduced two new parameters: \bar{c} and Λ_b .

Additional Parameters



In this process, we have introduced two new parameters: \bar{c} and Λ_b .

This changes the posterior we need to find:

Additional Parameters



In this process, we have introduced two new parameters: \bar{c} and Λ_b .

This changes the posterior we need to find:

$$\underbrace{\text{pr}(\vec{a}, \bar{c}^2, \Lambda_b | \vec{y}_{\text{exp}}, \mathbf{I})}_{\text{Total posterior}} \propto \underbrace{\text{pr}(\vec{y}_{\text{exp}} | \vec{a}, \Sigma, \mathbf{I})}_{\text{Likelihood for } \vec{a}} \underbrace{\text{pr}(\vec{a} | \mathbf{I})}_{\text{Prior for } \vec{a}}$$

Additional Parameters



In this process, we have introduced two new parameters: \bar{c} and Λ_b .

This changes the posterior we need to find:

$$\underbrace{\text{pr}(\vec{a}, \bar{c}^2, \Lambda_b | \vec{y}_{\text{exp}}, \mathbf{I})}_{\text{Total posterior}} \propto \underbrace{\text{pr}(\vec{y}_{\text{exp}} | \vec{a}, \Sigma, \mathbf{I})}_{\text{Likelihood for } \vec{a}} \underbrace{\text{pr}(\vec{a} | \mathbf{I})}_{\text{Prior for } \vec{a}} \underbrace{\text{pr}(\bar{c}^2 | \Lambda_b, \vec{a}, \mathbf{I})}_{\text{Posterior for } \bar{c}^2} \underbrace{\text{pr}(\Lambda_b | \vec{a}, \mathbf{I})}_{\text{Posterior for } \Lambda_b} .$$

Additional Parameters



In this process, we have introduced two new parameters: \bar{c} and Λ_b .

This changes the posterior we need to find:

$$\underbrace{\text{pr}(\vec{a}, \bar{c}^2, \Lambda_b | \vec{y}_{\text{exp}}, \mathbf{I})}_{\text{Total posterior}} \propto \underbrace{\text{pr}(\vec{y}_{\text{exp}} | \vec{a}, \Sigma, \mathbf{I})}_{\text{Likelihood for } \vec{a}} \underbrace{\text{pr}(\vec{a} | \mathbf{I})}_{\text{Prior for } \vec{a}} \underbrace{\text{pr}(\bar{c}^2 | \Lambda_b, \vec{a}, \mathbf{I})}_{\text{Posterior for } \bar{c}^2} \underbrace{\text{pr}(\Lambda_b | \vec{a}, \mathbf{I})}_{\text{Posterior for } \Lambda_b} .$$

We can find a closed form of $\text{pr}(\bar{c}^2 | \Lambda_b, \vec{a}, \mathbf{I})$ and $\text{pr}(\Lambda_b | \vec{a}, \mathbf{I})$.

Posterior for \bar{c}





Posterior for \bar{c}

Since we had $c_n | \bar{c} \sim \mathcal{N}(0, \bar{c}^2)$, where \bar{c}^2 is a population variance, we make the standard choice of prior for an unknown variance:

$$\bar{c}^2 \sim \chi^{-2}(\nu_0, \tau_0^2)$$



Posterior for \bar{c}

Since we had $c_n | \bar{c} \sim \mathcal{N}(0, \bar{c}^2)$, where \bar{c}^2 is a population variance, we make the standard choice of prior for an unknown variance:

$$\bar{c}^2 \sim \chi^{-2}(\nu_0, \tau_0^2)$$

This yields a conjugate posterior

$$\text{pr}(\bar{c}^2 | \mathbf{I}) \sim \chi^{-2}(\nu_0, \tau_0^2) \iff \text{pr}(\bar{c}^2 | \vec{\mathbf{a}}, \Lambda_{\mathbf{b}}, \mathbf{I}) \sim \chi^{-2}(\nu, \tau^2(\vec{\mathbf{a}}, \Lambda_{\mathbf{b}})).$$



Posterior for \bar{c}

Since we had $c_n | \bar{c} \sim \mathcal{N}(0, \bar{c}^2)$, where \bar{c}^2 is a population variance, we make the standard choice of prior for an unknown variance:

$$\bar{c}^2 \sim \chi^{-2}(\nu_0, \tau_0^2)$$

This yields a conjugate posterior

$$\text{pr}(\bar{c}^2 | \mathbf{I}) \sim \chi^{-2}(\nu_0, \tau_0^2) \iff \text{pr}(\bar{c}^2 | \vec{\mathbf{a}}, \Lambda_{\mathbf{b}}, \mathbf{I}) \sim \chi^{-2}(\nu, \tau^2(\vec{\mathbf{a}}, \Lambda_{\mathbf{b}})).$$

Where we have hyperparameters:



Posterior for \bar{c}

Since we had $c_n | \bar{c} \sim \mathcal{N}(0, \bar{c}^2)$, where \bar{c}^2 is a population variance, we make the standard choice of prior for an unknown variance:

$$\bar{c}^2 \sim \chi^{-2}(\nu_0, \tau_0^2)$$

This yields a conjugate posterior

$$\text{pr}(\bar{c}^2 | \text{I}) \sim \chi^{-2}(\nu_0, \tau_0^2) \iff \text{pr}(\bar{c}^2 | \vec{a}, \Lambda_b, \text{I}) \sim \chi^{-2}(\nu, \tau^2(\vec{a}, \Lambda_b)).$$

Where we have hyperparameters:

$$\nu = \nu_0 + N_{\text{obs}} n_{\text{orders}}, \text{ degrees of freedom}$$



Posterior for \bar{c}

Since we had $c_n | \bar{c} \sim \mathcal{N}(0, \bar{c}^2)$, where \bar{c}^2 is a population variance, we make the standard choice of prior for an unknown variance:

$$\bar{c}^2 \sim \chi^{-2}(\nu_0, \tau_0^2)$$

This yields a conjugate posterior

$$\text{pr}(\bar{c}^2 | \mathbf{I}) \sim \chi^{-2}(\nu_0, \tau_0^2) \iff \text{pr}(\bar{c}^2 | \vec{a}, \Lambda_b, \mathbf{I}) \sim \chi^{-2}(\nu, \tau^2(\vec{a}, \Lambda_b)).$$

Where we have hyperparameters:

$$\nu = \nu_0 + N_{\text{obs}} n_{\text{orders}}, \text{ degrees of freedom}$$

$$\tau^2(\vec{a}, \Lambda_b) = \frac{1}{\nu} \left(\nu_0 \tau_0 + \sum_{i,n} c_{n,i}^2(\vec{a}, \Lambda_b) \right), \text{ scale}$$



Posterior for \bar{c}

Since we had $c_n | \bar{c} \sim \mathcal{N}(0, \bar{c}^2)$, where \bar{c}^2 is a population variance, we make the standard choice of prior for an unknown variance:

$$\bar{c}^2 \sim \chi^{-2}(\nu_0, \tau_0^2)$$

This yields a conjugate posterior

$$\text{pr}(\bar{c}^2 | \mathbf{I}) \sim \chi^{-2}(\nu_0, \tau_0^2) \iff \text{pr}(\bar{c}^2 | \vec{a}, \Lambda_b, \mathbf{I}) \sim \chi^{-2}(\nu, \tau^2(\vec{a}, \Lambda_b)).$$

Where we have hyperparameters:

$$\nu = \nu_0 + N_{\text{obs}} n_{\text{orders}}, \text{ degrees of freedom}$$
$$\tau^2(\vec{a}, \Lambda_b) = \frac{1}{\nu} \left(\nu_0 \tau_0 + \sum_{i,n} c_{n,i}^2(\vec{a}, \Lambda_b) \right), \text{ scale}$$
$$c_{n,i} = \frac{y_i^n - y_i^{(n-1)}}{y_{\text{ref},i} Q_i^n}$$

Posterior for Λ_b



Posterior for Λ_b



Our posterior for the breakdown scale also uses these hyperparameters:



Posterior for Λ_b

Our posterior for the breakdown scale also uses these hyperparameters:

$$\text{pr}(\Lambda_b | \vec{a}, \mathbf{I}) \propto \frac{\text{pr}(\Lambda_b | \mathbf{I})}{\tau^\nu \prod_{n,i} \left(\frac{p_i}{\Lambda_b} \right)^n}$$



Posterior for Λ_b

Our posterior for the breakdown scale also uses these hyperparameters:

$$\text{pr}(\Lambda_b | \vec{a}, \mathbf{I}) \propto \frac{\text{pr}(\Lambda_b | \mathbf{I})}{\tau^\nu \prod_{n,i} \left(\frac{p_i}{\Lambda_b} \right)^n}$$

This posterior needs to be numerically normalized as the normalization constant is dependent on \vec{a} .



Posterior for Λ_b

Our posterior for the breakdown scale also uses these hyperparameters:

$$\text{pr}(\Lambda_b | \vec{a}, \mathbf{I}) \propto \frac{\text{pr}(\Lambda_b | \mathbf{I})}{\tau^\nu \prod_{n,i} \left(\frac{p_i}{\Lambda_b} \right)^n}$$

This posterior needs to be numerically normalized as the normalization constant is dependent on \vec{a} .

With all our components, we can estimate our parameters.

Outline

- Motivation/Bayesian EFT Model Calibration
- BUQEYE Formalism
- **Interaction Choice**
- Results
- Questions

Interaction Choice



Interaction Choice



Hidden in the I in our distributions is the choice of interaction.

Interaction Choice



Hidden in the I in our distributions is the choice of interaction.

This includes:

Interaction Choice



Hidden in the I in our distributions is the choice of interaction.

This includes:

- Degrees of freedom

Interaction Choice



Hidden in the I in our distributions is the choice of interaction.

This includes:

- Degrees of freedom
- Power counting

Interaction Choice



Hidden in the I in our distributions is the choice of interaction.

This includes:

- Degrees of freedom
- Power counting
- Representation

Interaction Choice



Hidden in the I in our distributions is the choice of interaction.

This includes:

- Degrees of freedom
- Power counting
- Representation
- Regularization scheme

Pionless EFT



Pionless EFT



We are working in a “Weinberg-ized” pionless EFT.

Pionless EFT

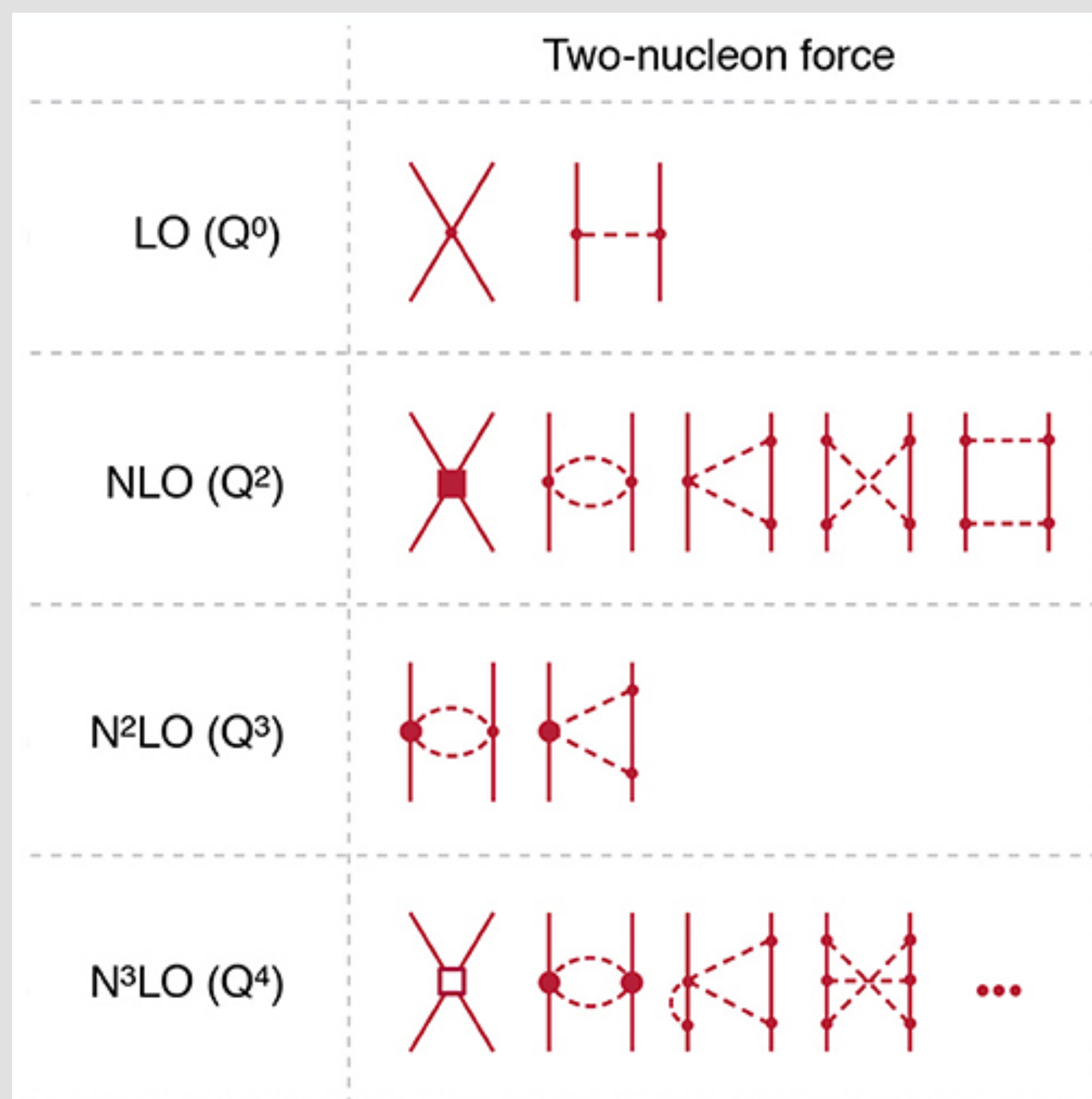


We are working in a “Weinberg-ized” pionless EFT.

Pionless EFT



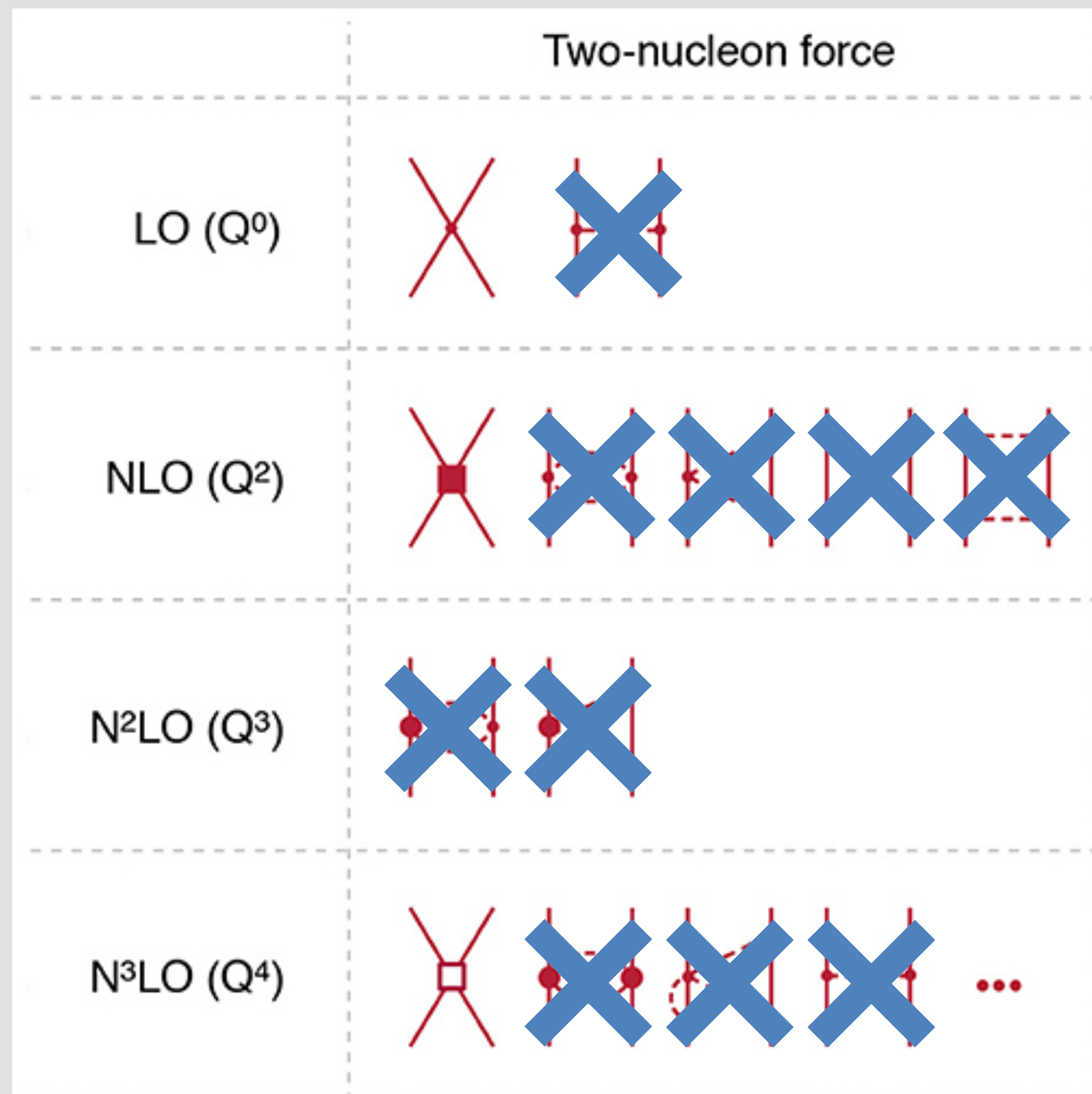
We are working in a “Weinberg-ized” pionless EFT.





Pionless EFT

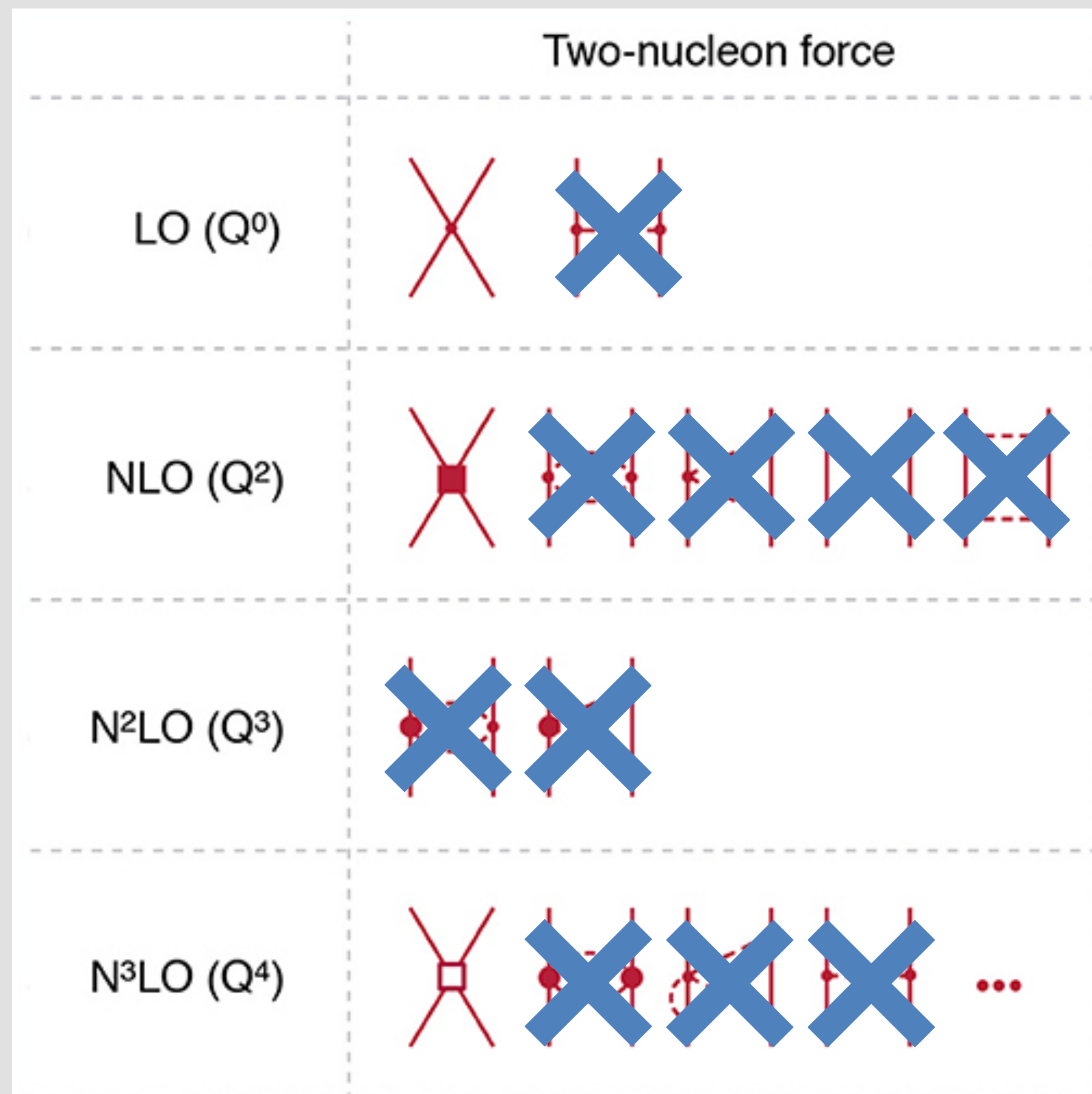
We are working in a “Weinberg-ized” pionless EFT.





Pionless EFT

We are working in a “Weinberg-ized” pionless EFT.

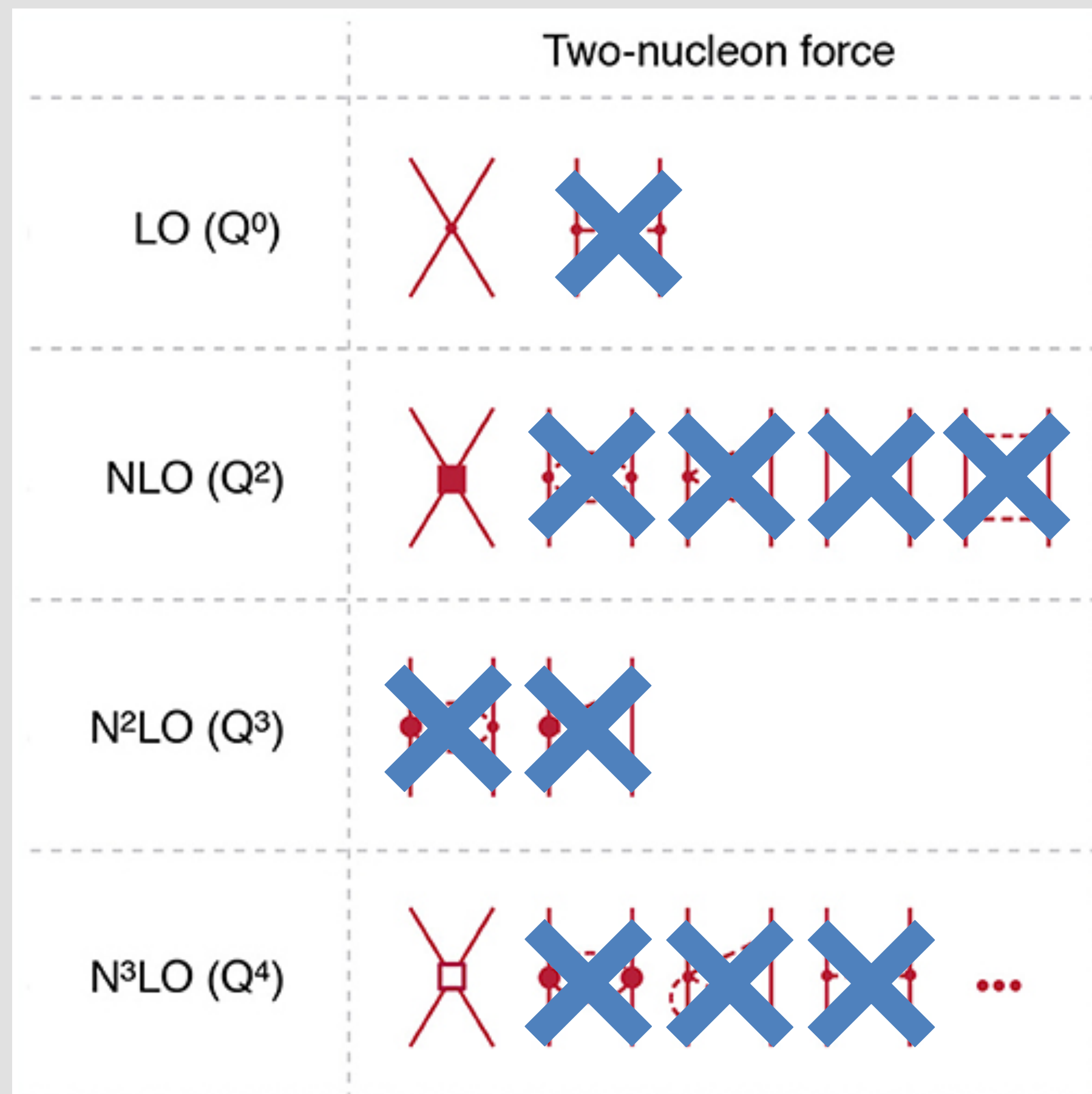


$$y_{\text{th}}(x) = y_{\text{ref}}(x) \sum_{n=0}^{\infty} c_{2n}(x) Q^{2n}(x)$$



Pionless EFT

We are working in a “Weinberg-ized” pionless EFT.



$$y_{\text{th}}(x) = \underbrace{y_{\text{ref}}(x)}_{y_{\text{exp}}(x)} \sum_{n=0}^{\infty} c_{2n}(x) Q^{2n}(x)$$

Pionless EFT



We are working in a “Weinberg-ized” pionless EFT.

Our interaction takes the form:

$$y_{\text{th}}(x) = \frac{y_{\text{exp}}(x)}{y_{\text{ref}}(x)} \sum_{n=0}^{\infty} c_{2n}(x) Q^{2n}(x)$$



Pionless EFT

We are working in a “Weinberg-ized” pionless EFT.

Our interaction takes the form:

$$y_{\text{th}}(x) = \frac{y_{\text{ref}}(x)}{y_{\text{exp}}(x)} \sum_{n=0}^{\infty} c_{2n}(x) Q^{2n}(x)$$

$$v_{\text{LO}} = C_S + C_T \sigma_1 \cdot \sigma_2$$

Pionless EFT



We are working in a “Weinberg-ized” pionless EFT.

Our interaction takes the form:

$$y_{\text{th}}(x) = \frac{y_{\text{ref}}(x)}{y_{\text{exp}}(x)} \sum_{n=0}^{\infty} c_{2n}(x) Q^{2n}(x)$$

$$v_{\text{LO}} = C_S + C_T \sigma_1 \cdot \sigma_2$$

$$v_{\text{NLO}}^{\text{CI}}(\vec{k}, \vec{K}) = C_1 k^2 + C_2 k^2 \sigma_1 \cdot \sigma_2 + C_3 S_{12}(k) + C_4 k^2 \tau_1 \cdot \tau_2 \\ + iC_5 \vec{S} \cdot (\vec{K} \times \vec{k}) + C_6 k^2 \tau_1 \cdot \tau_2 \sigma_1 \cdot \sigma_2 + C_7 S_{12}(k) \tau_2 \cdot \tau_2$$

Pionless EFT



We are working in a “Weinberg-ized” pionless EFT.

Our interaction takes the form:

$$y_{\text{th}}(x) = \frac{y_{\text{ref}}(x)}{y_{\text{exp}}(x)} \sum_{n=0}^{\infty} c_{2n}(x) Q^{2n}(x)$$

$$v_{\text{LO}} = C_S + C_T \sigma_1 \cdot \sigma_2$$

$$\begin{aligned} v_{\text{NLO}}^{\text{CI}}(\vec{k}, \vec{K}) &= C_1 k^2 + C_2 k^2 \sigma_1 \cdot \sigma_2 + C_3 S_{12}(k) + C_4 k^2 \tau_1 \cdot \tau_2 \\ &+ iC_5 \vec{S} \cdot (\vec{K} \times \vec{k}) + C_6 k^2 \tau_1 \cdot \tau_2 \sigma_1 \cdot \sigma_2 + C_7 S_{12}(k) \tau_2 \cdot \tau_2 \\ v_{\text{NLO}}^{\text{CD}} &= C_0^{\text{IT}} T_{12} + C_0^{\text{IV}} (\tau_{1z} + \tau_{2z}) \end{aligned}$$

Regularization



Regularization



To use these interactions, they must be regularized in some fashion and must be local in coordinate space (for QMC).

Regularization



To use these interactions, they must be regularized in some fashion and must be local in coordinate space (for QMC).

We employ a Gaussian cutoff in coordinate space, which smears δ -functions upon Fourier transformation

$$f(r) = \frac{1}{\pi^{3/2} R_s^3} e^{-\left(\frac{r}{R_s}\right)^2}$$

Regularization



To use these interactions, they must be regularized in some fashion and must be local in coordinate space (for QMC).

We employ a Gaussian cutoff in coordinate space, which smears δ -functions upon Fourier transformation

$$f(r) = \frac{1}{\pi^{3/2} R_s^3} e^{-\left(\frac{r}{R_s}\right)^2}$$

We choose $R_s \in [1.5, 2.0, 2.5]$ fm which are $\sim \frac{400}{R_s}$ MeV in momentum space.

Parameter Estimation Algorithm



Parameter Estimation Algorithm



To estimate all of these parameters, we need data to calibrate to:

Parameter Estimation Algorithm



To estimate all of these parameters, we need data to calibrate to:

Our choice of data is the pp and np Granada database (251 differential cross sections, 133 total cross sections, 4 polarized cross sections) up to 5 MeV + deuteron binding energy + nn scattering length.



Parameter Estimation Algorithm

To estimate all of these parameters, we need data to calibrate to:

Our choice of data is the pp and np Granada database (251 differential cross sections, 133 total cross sections, 4 polarized cross sections) up to 5 MeV + deuteron binding energy + nn scattering length.

We then use Markov Chain Monte Carlo (MCMC) to sample the posteriors at LO (Q^0), NLO (Q^2), and N3LO (Q^4), allowing for the order-by-order convergence analysis for LO \rightarrow NLO and NLO \rightarrow N3LO to estimate \bar{c} and Λ_b .



Prior Choices

- $\text{pr}(\vec{a} | \text{I}) \sim \mathcal{N} \left(\vec{a}_{\text{p.s.}}^{\text{MAP}}, \overline{10^2} \right)$
- $\text{pr}(\Lambda_b | \text{I}) \sim \mathcal{N} (500 \text{ MeV}, 1000^2 \text{ MeV}^2)$
- $\text{pr}(\bar{c}^2 | \text{I}) \sim \chi^{-2}(\nu_0 = 1.5, \tau_0^2 = 1.5^2)$
- $r(x_i, x_j; \vec{l}) = e^{|p_i - p_j|/2l_p} e^{|\theta_i - \theta_j|/2l_\theta} \delta_{\text{type}_i, \text{type}_j}, \quad l_p = 0.3 \text{ MeV}, l_\theta = 20^\circ$
- $P_{\text{soft}} = \begin{cases} p_d \sim 45 \text{ MeV}/c, & \text{for } np \text{ scattering} \\ 1/{}^1a_{pp} \sim 25 \text{ MeV}, & \text{for } pp \text{ scattering.} \end{cases}$

Outline

- Motivation/Bayesian EFT Model Calibration
- BUQEYE Formalism
- Interaction Choice
- **Results**
- Questions

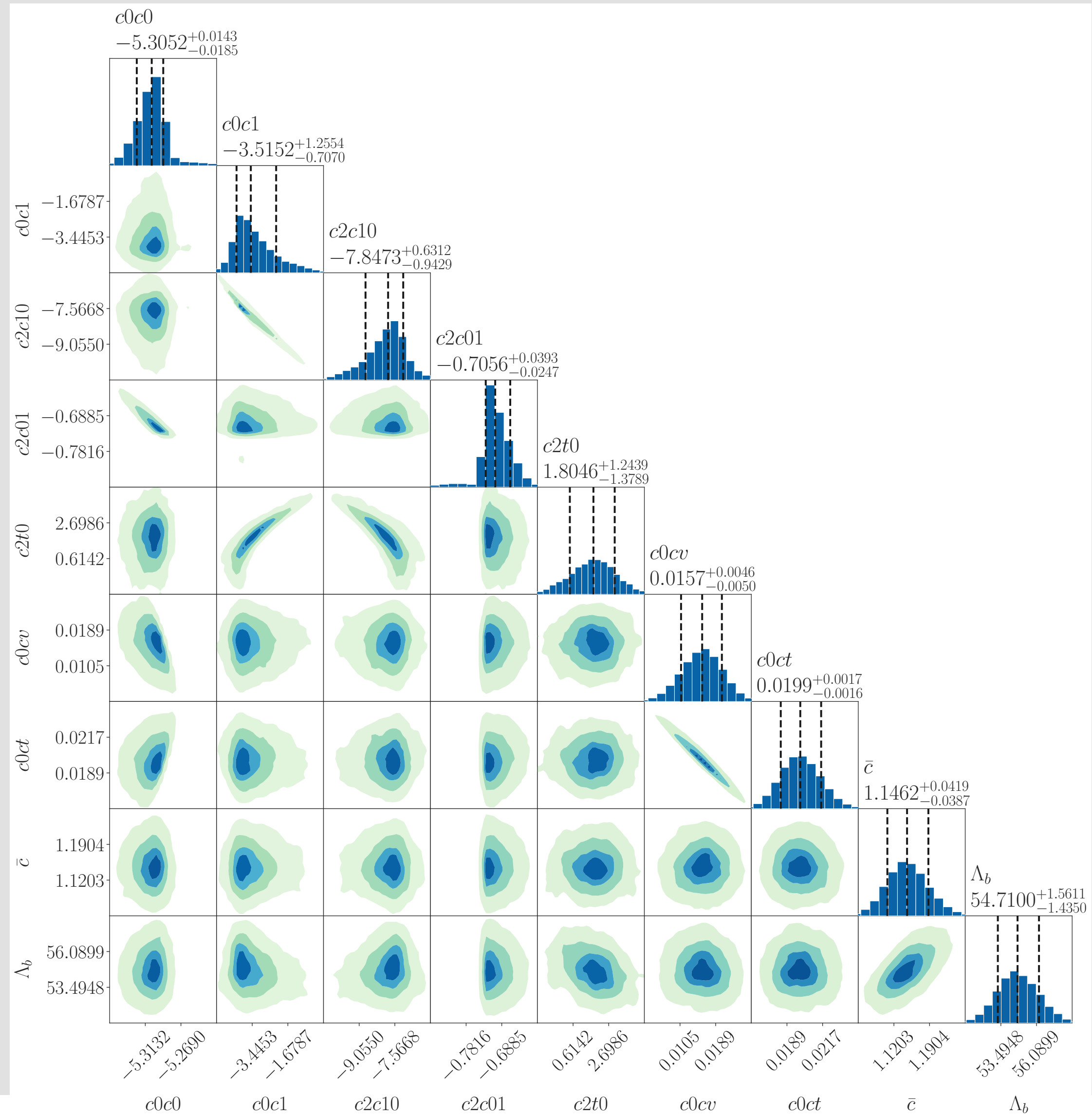
NLO Posterior



[Figures:](#)



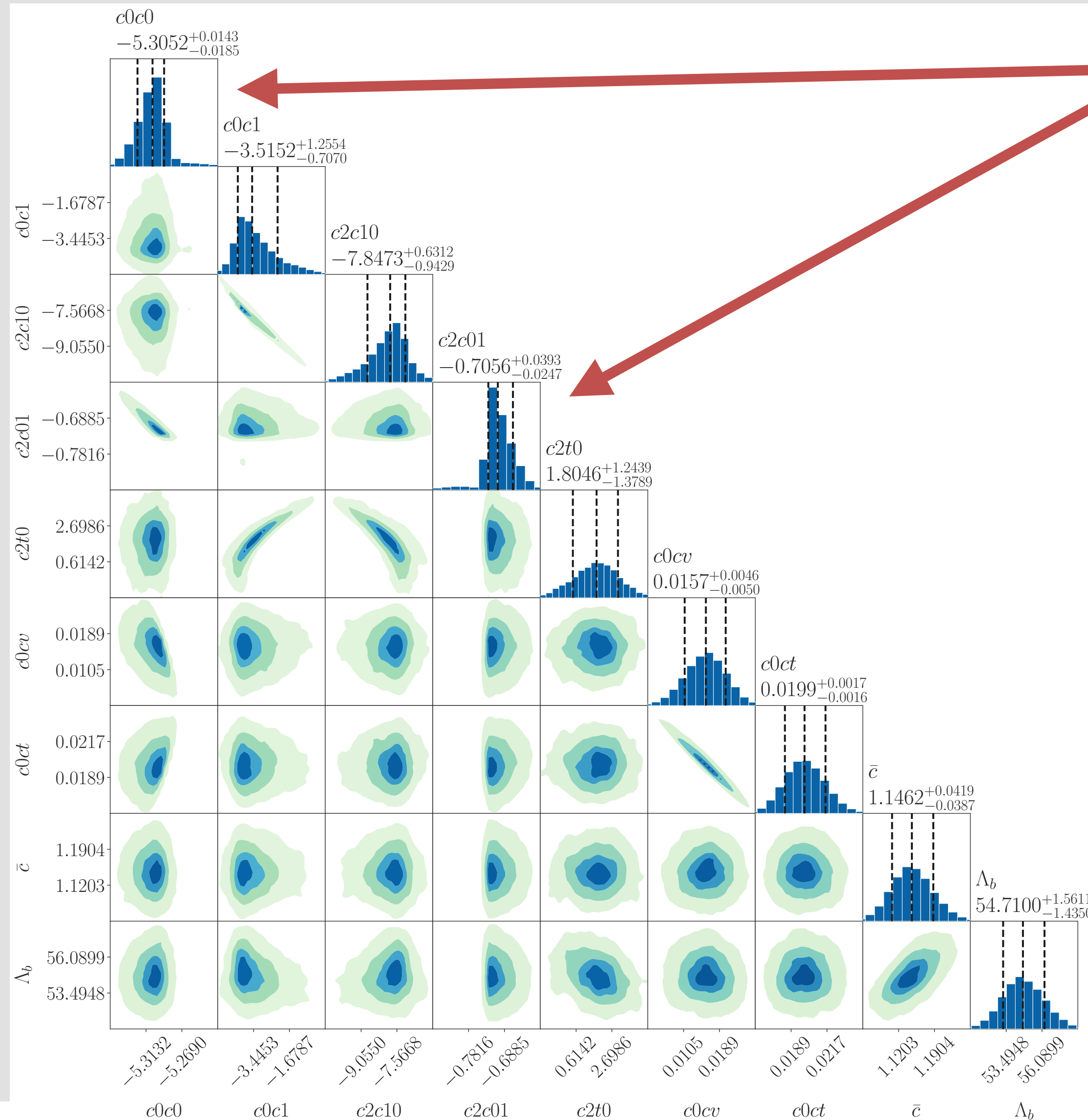
NLO Posterior



Figures:



NLO Posterior

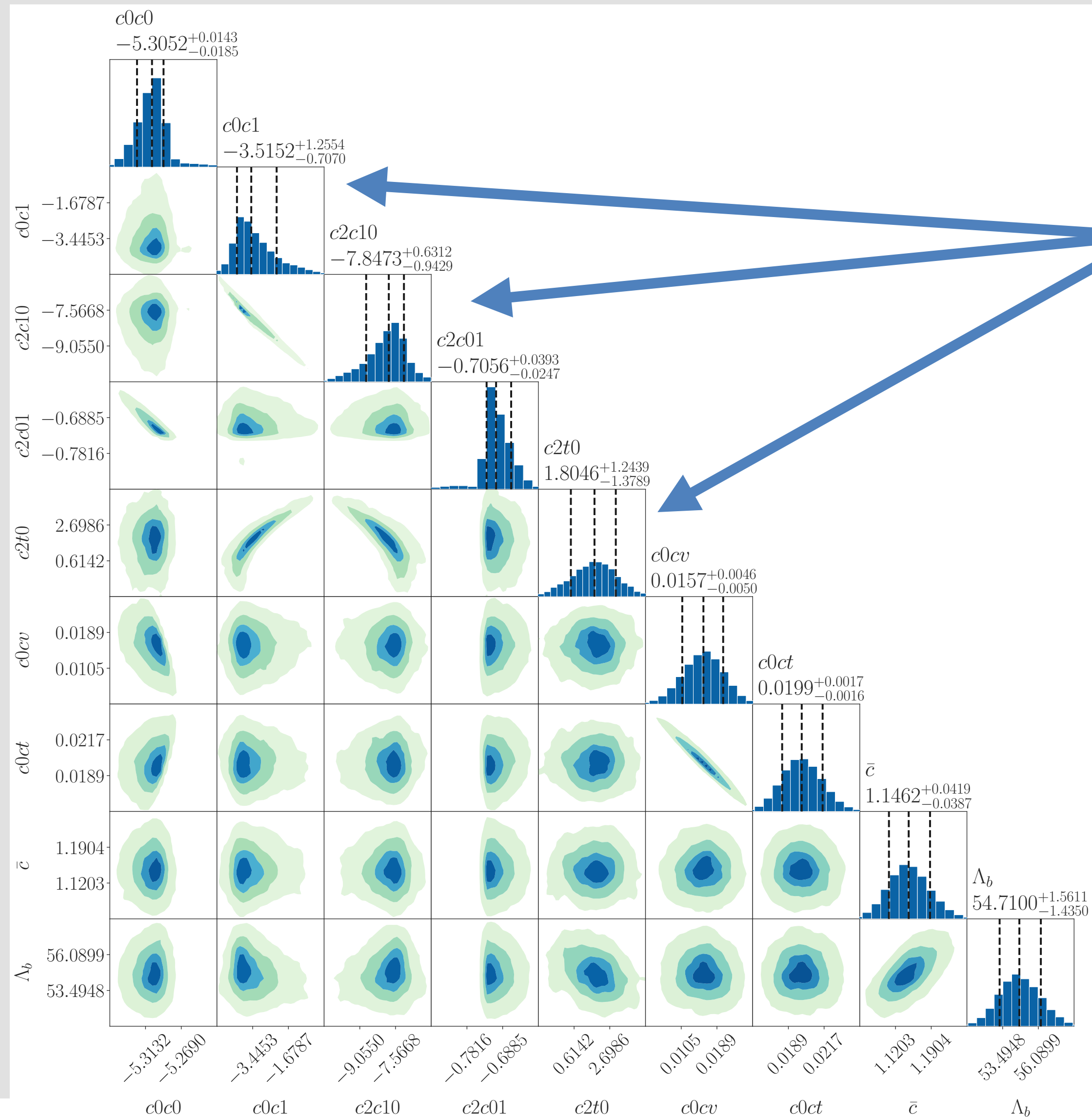


$$(S, T) = (0, 1)$$

Figures:



NLO Posterior

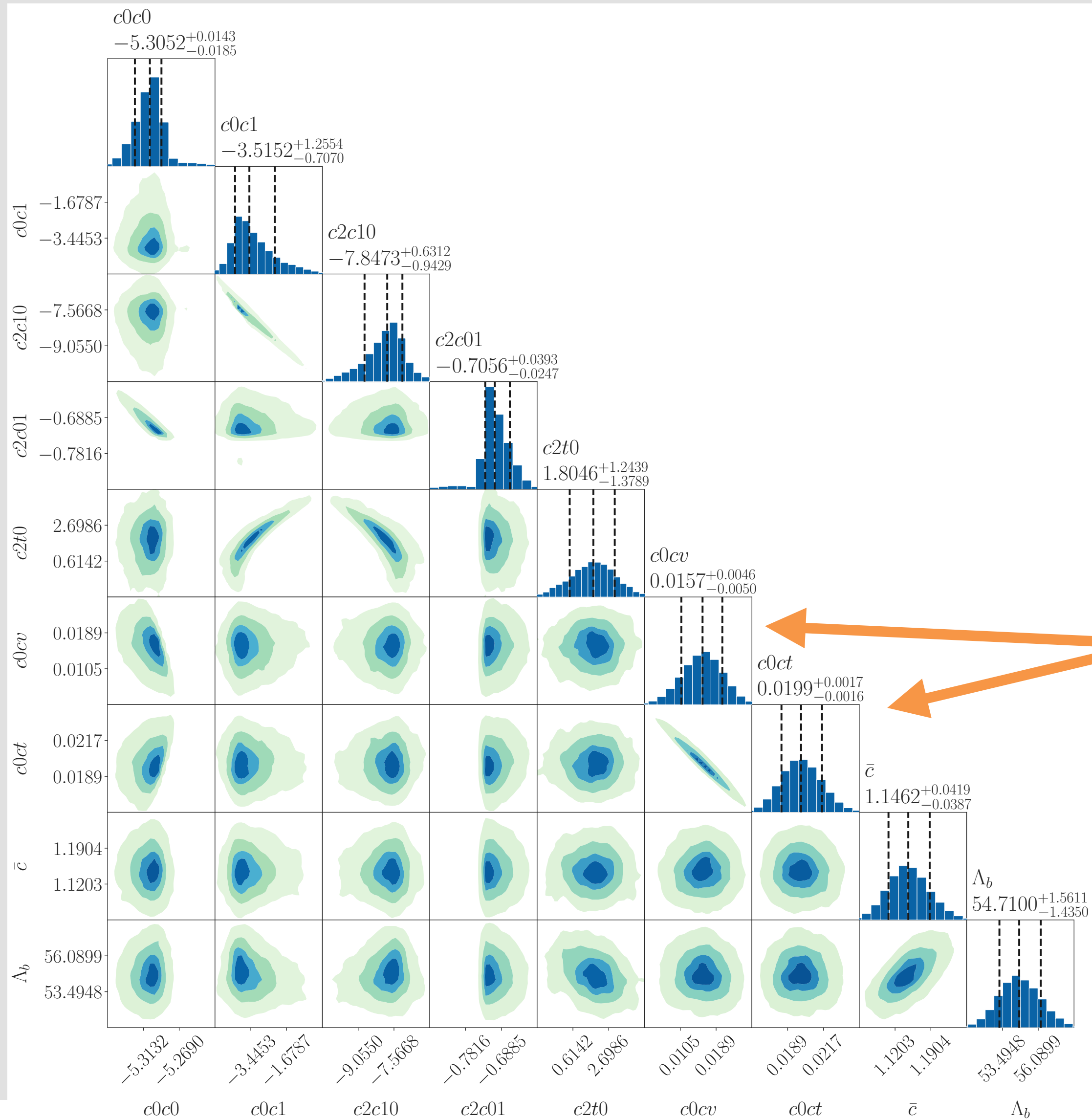


$$(S, T) = (1, 0)$$

Figures:



NLO Posterior

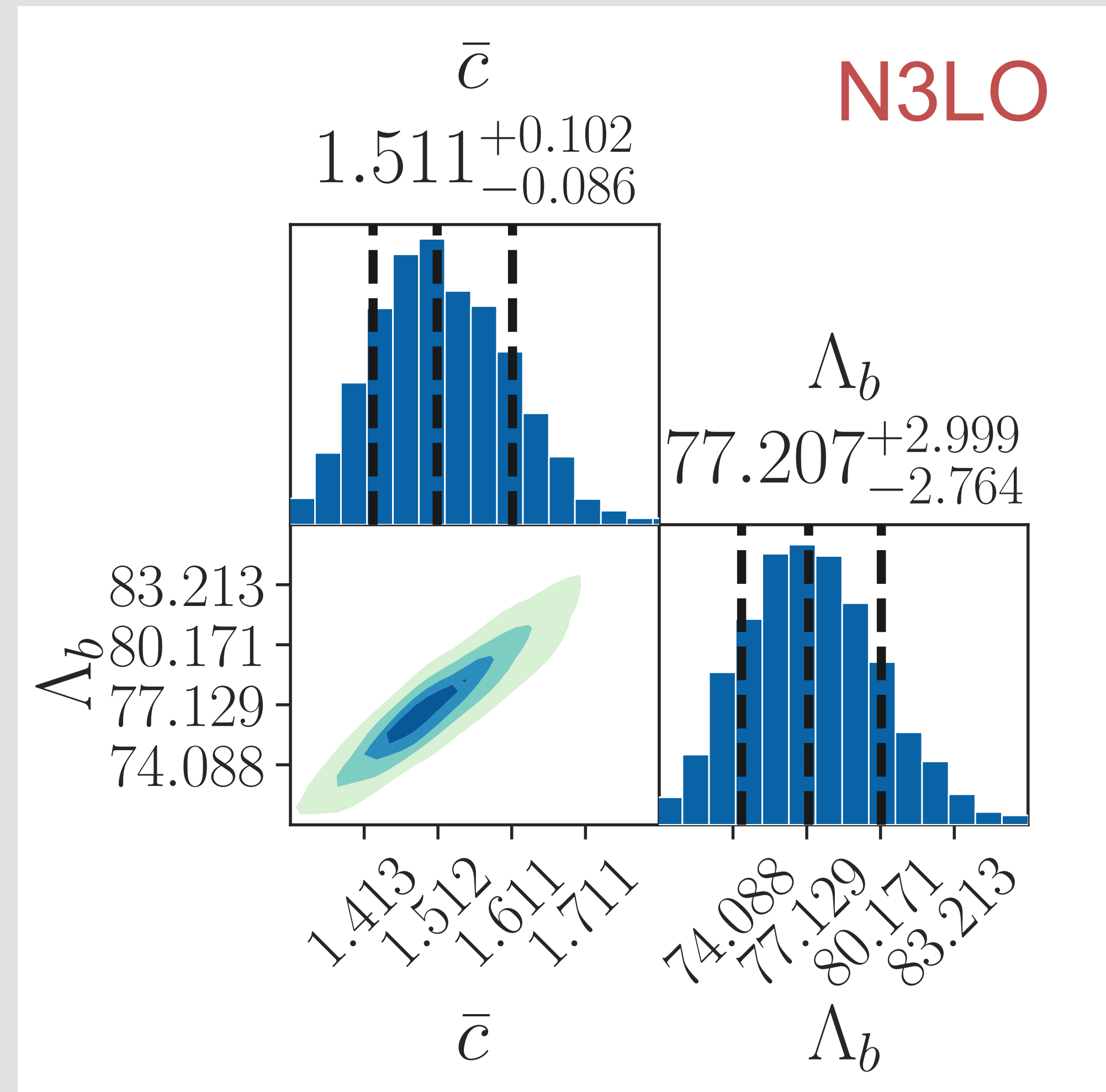
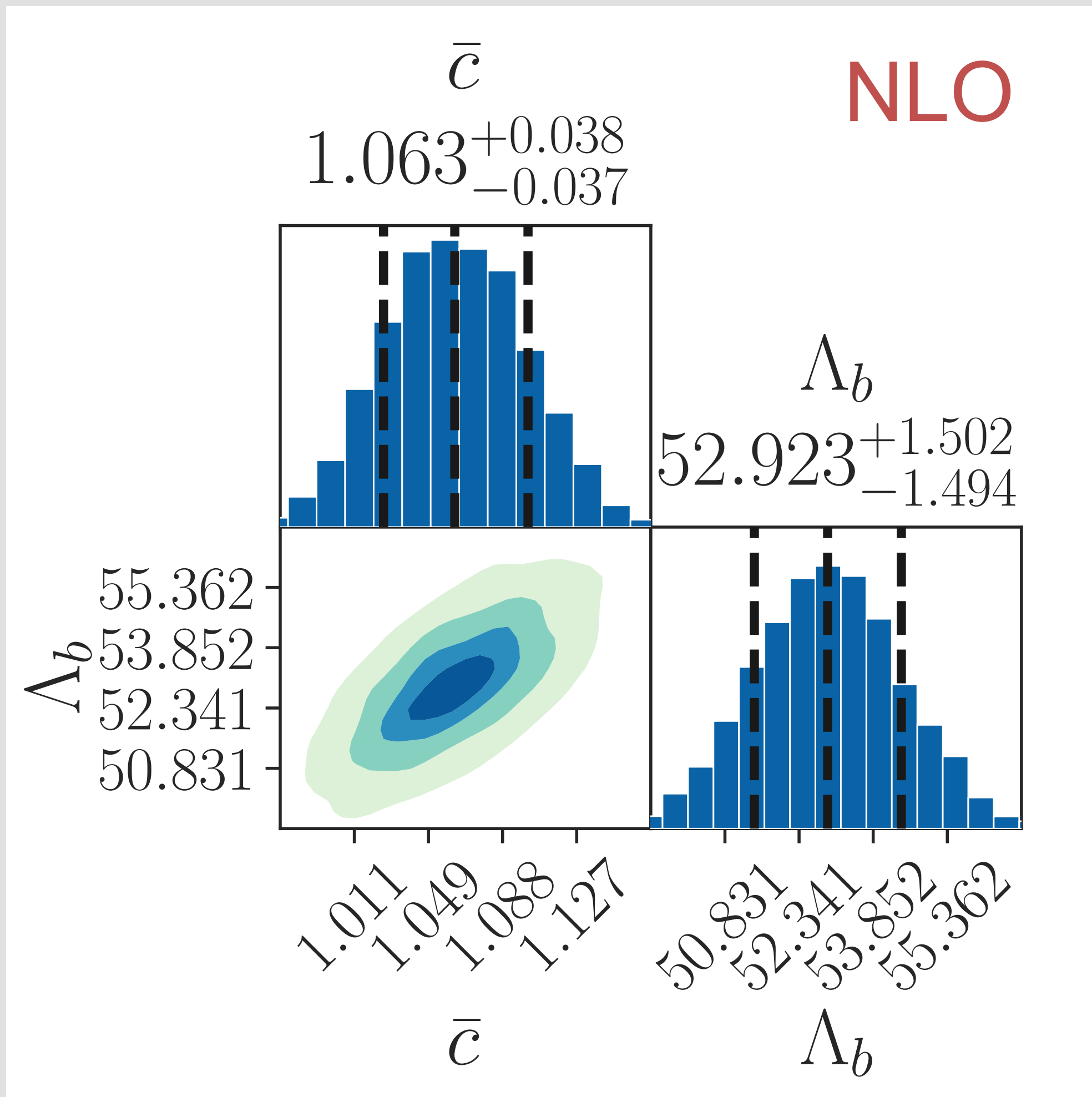


C^{CD}

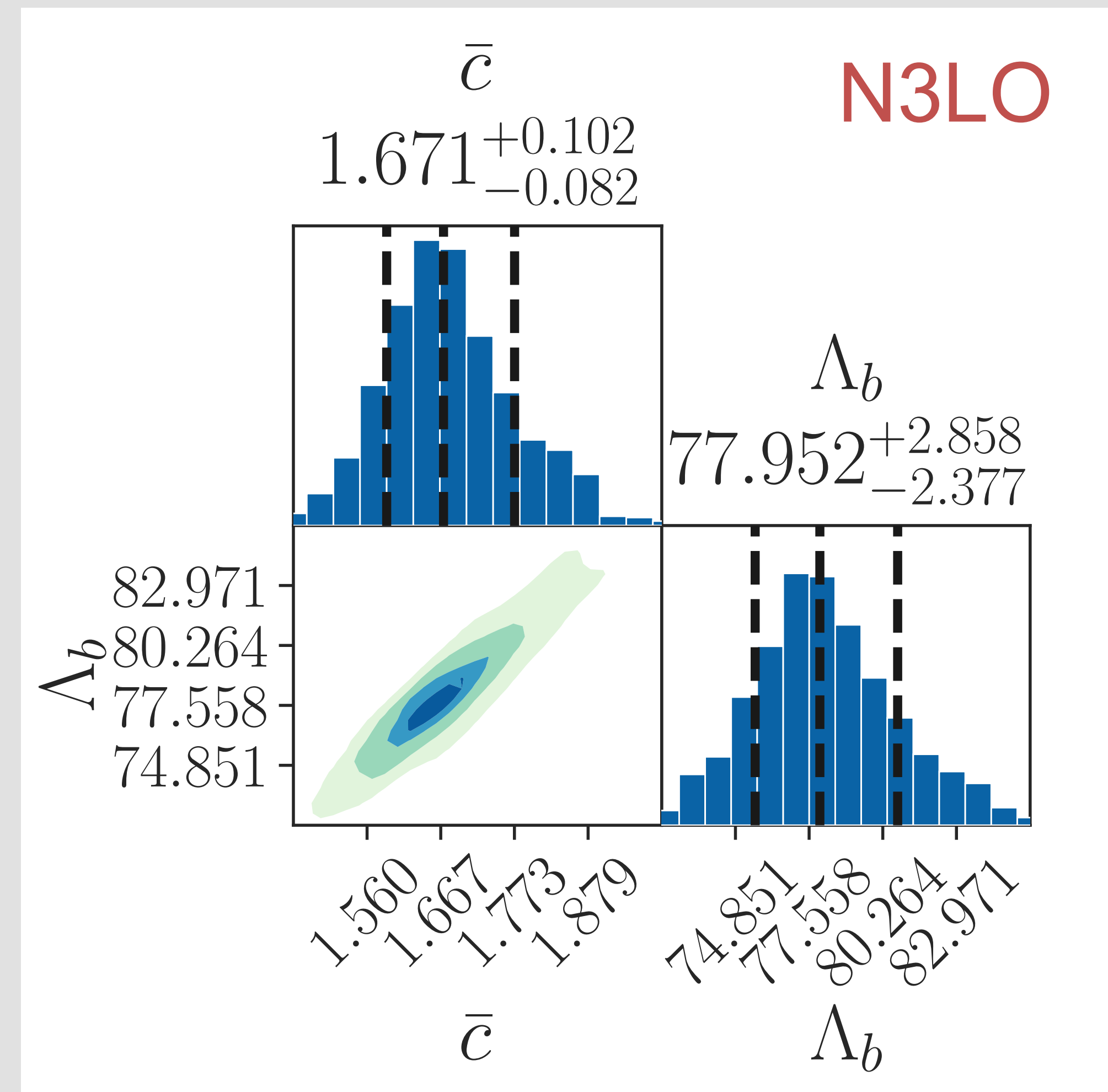
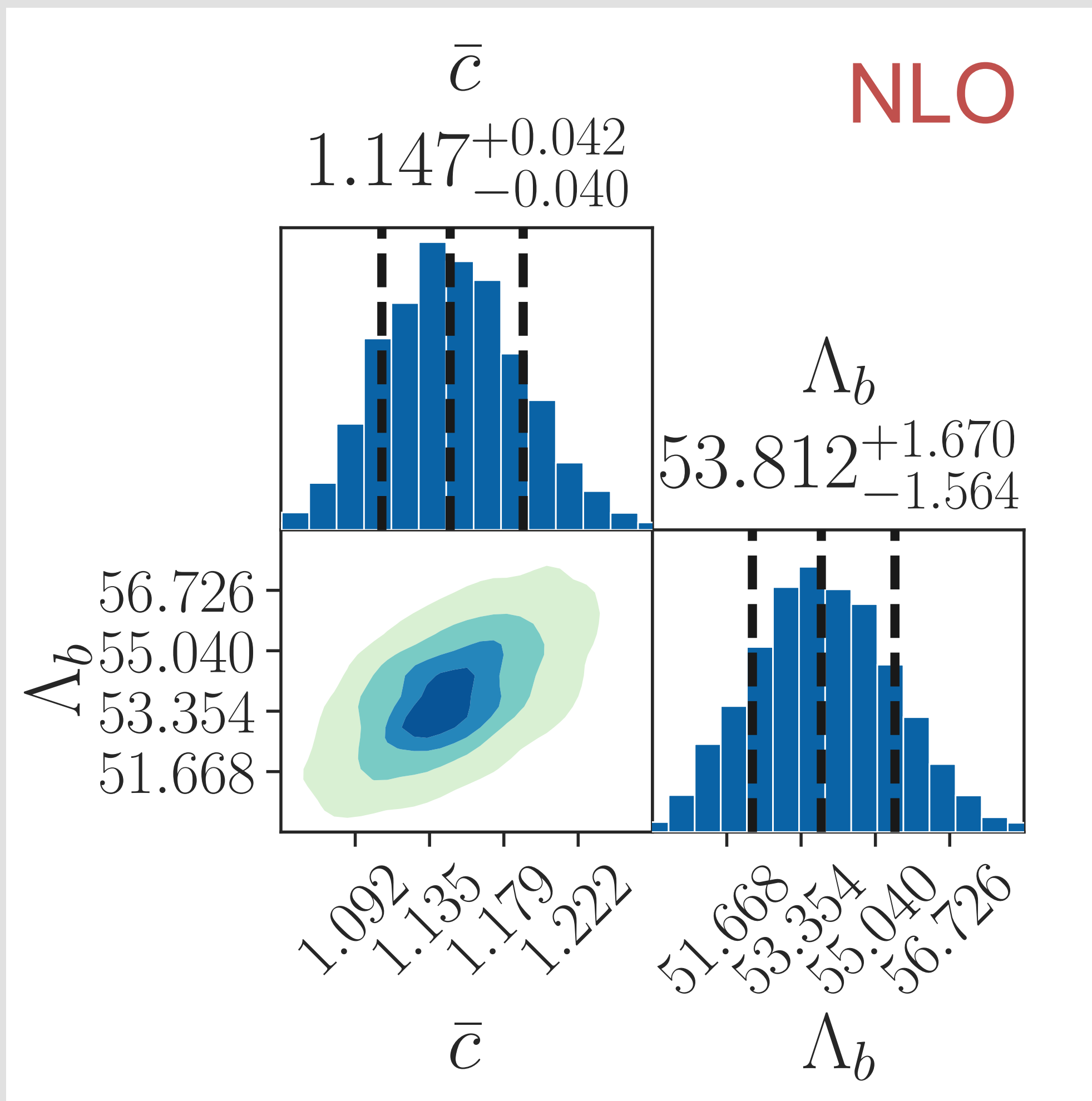
Figures:



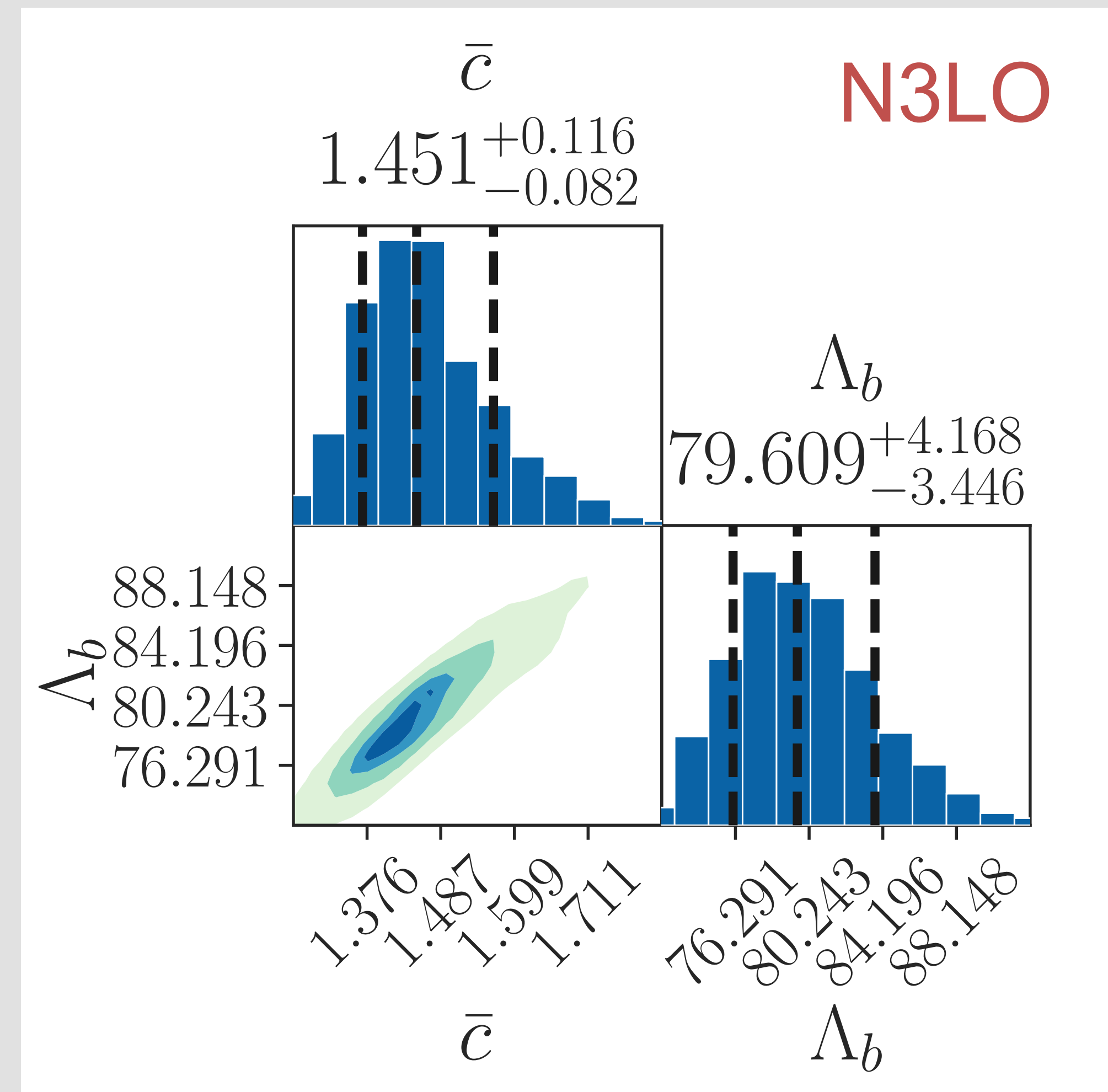
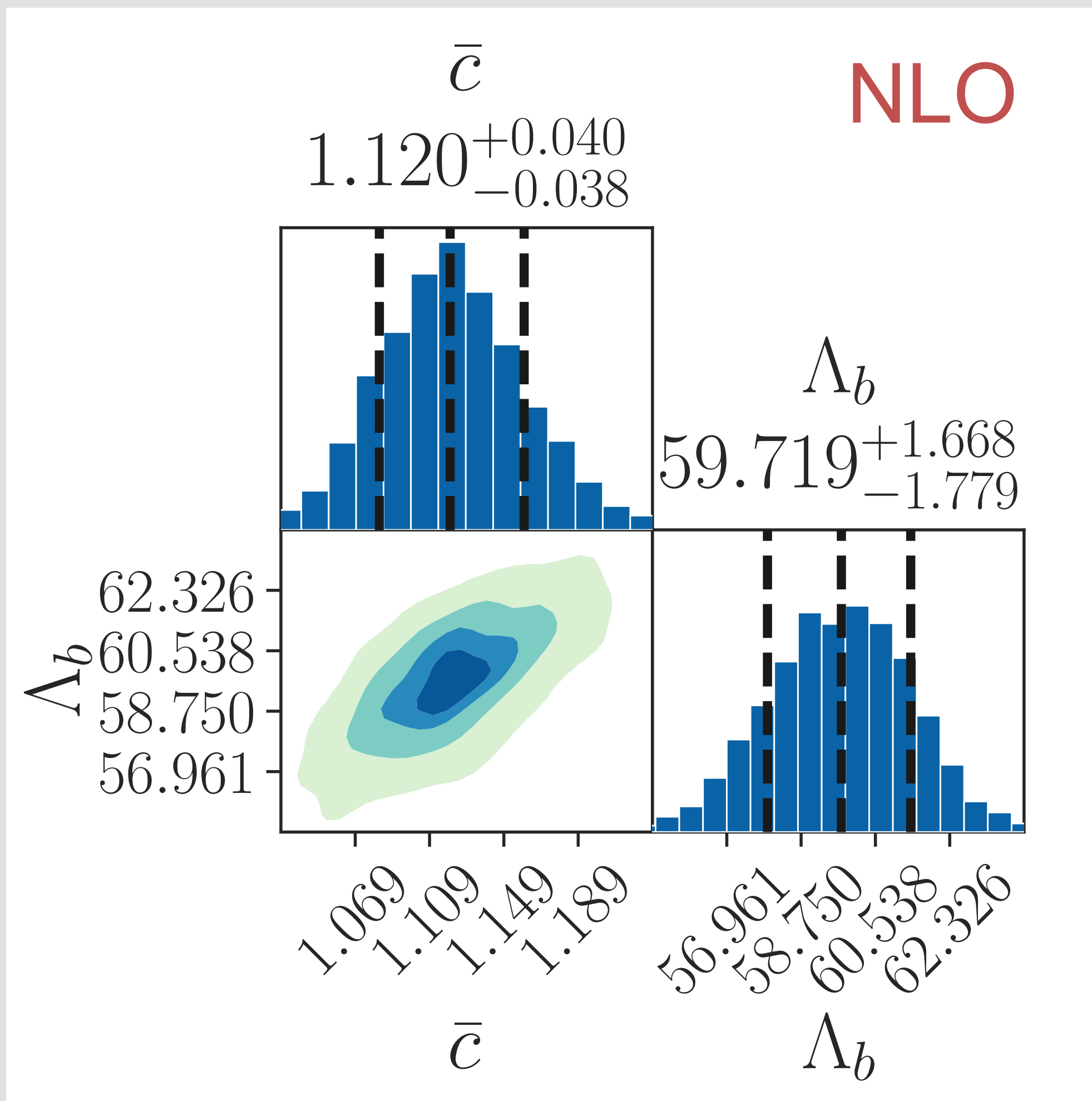
2.5 fm \bar{c} and Λ_b Posteriors



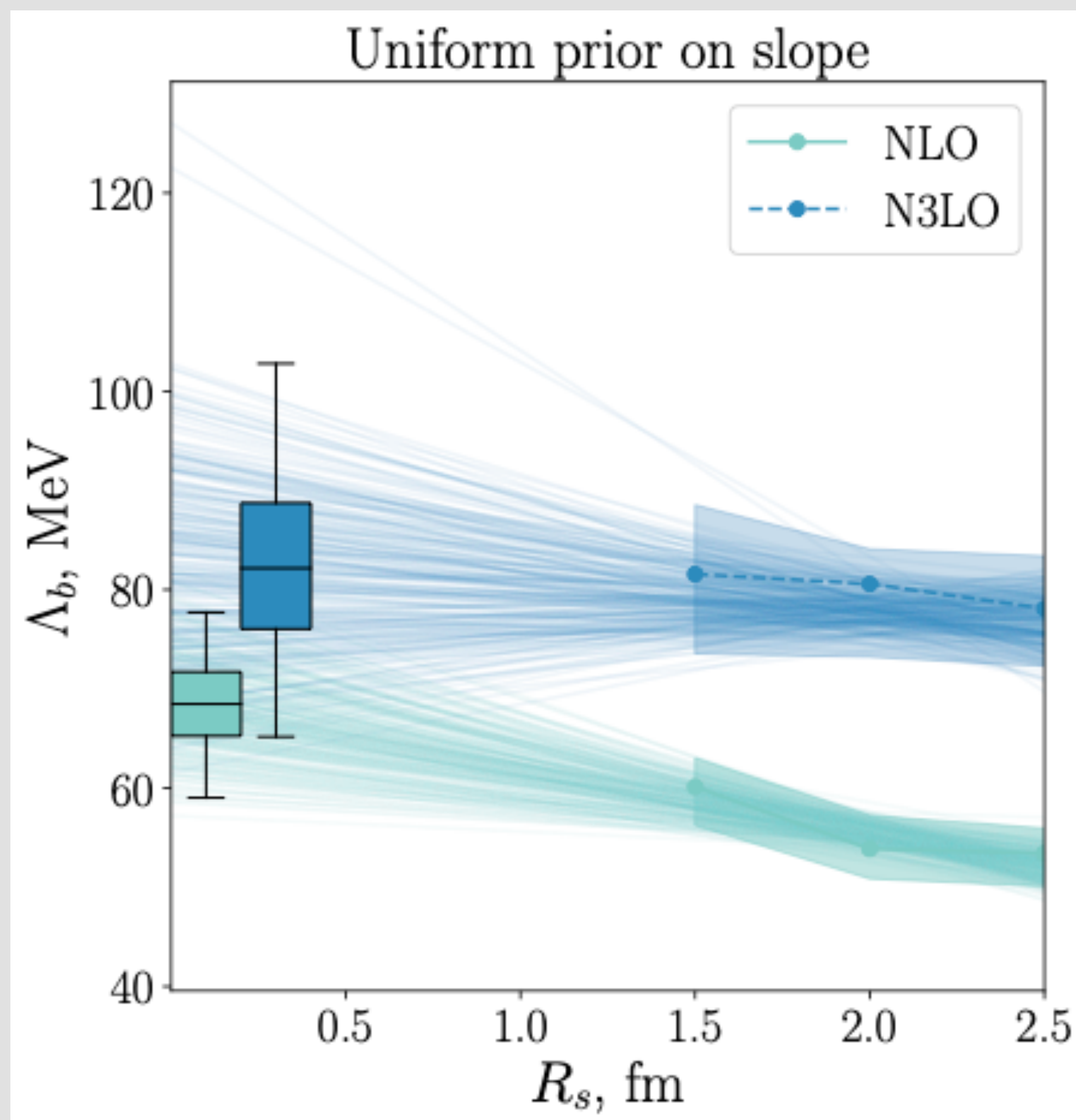
2.0 fm \bar{c} and Λ_b Posteriors



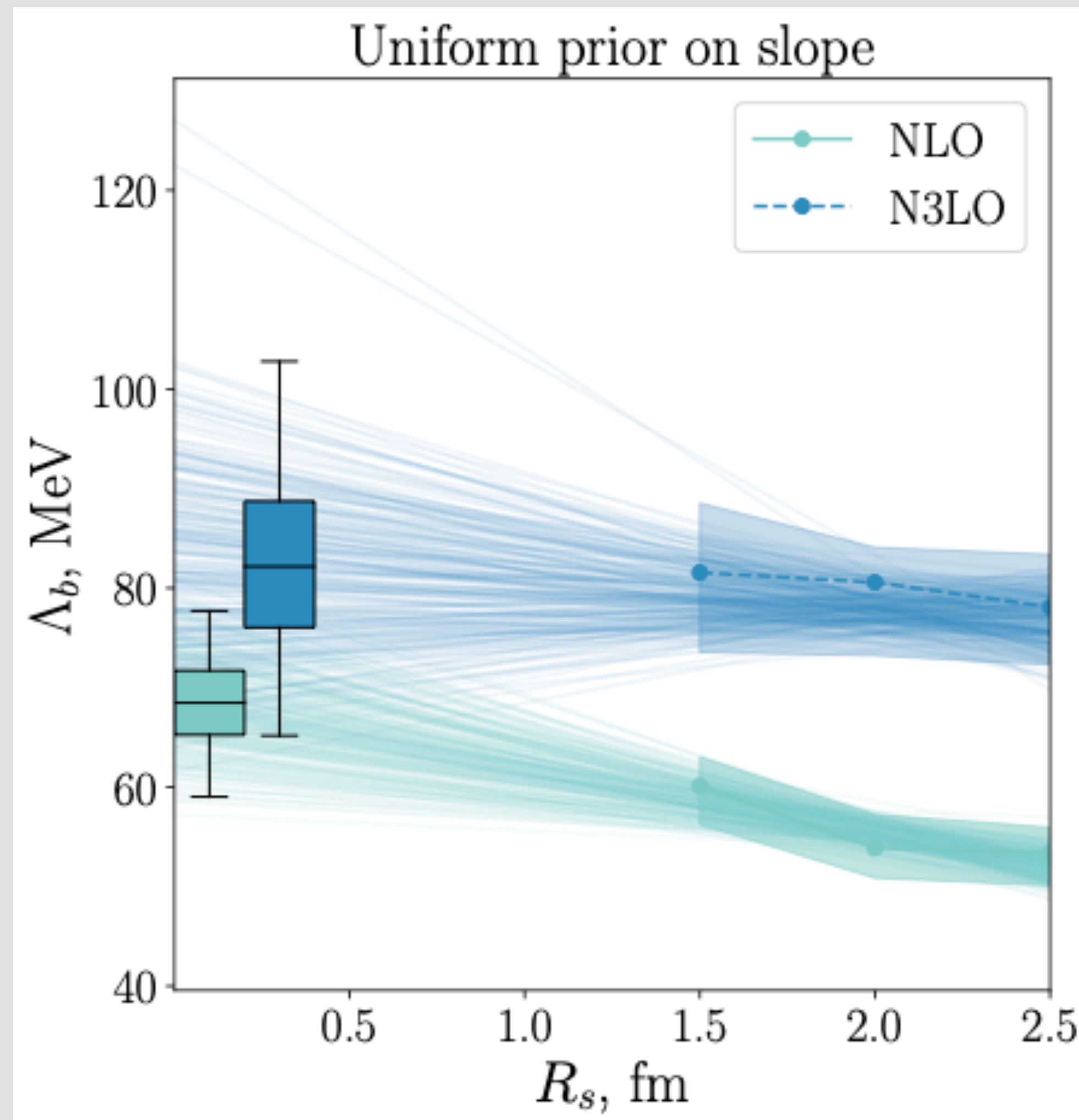
1.5 fm \bar{c} and Λ_b Posteriors



Extrapolation to Remove Artifacts

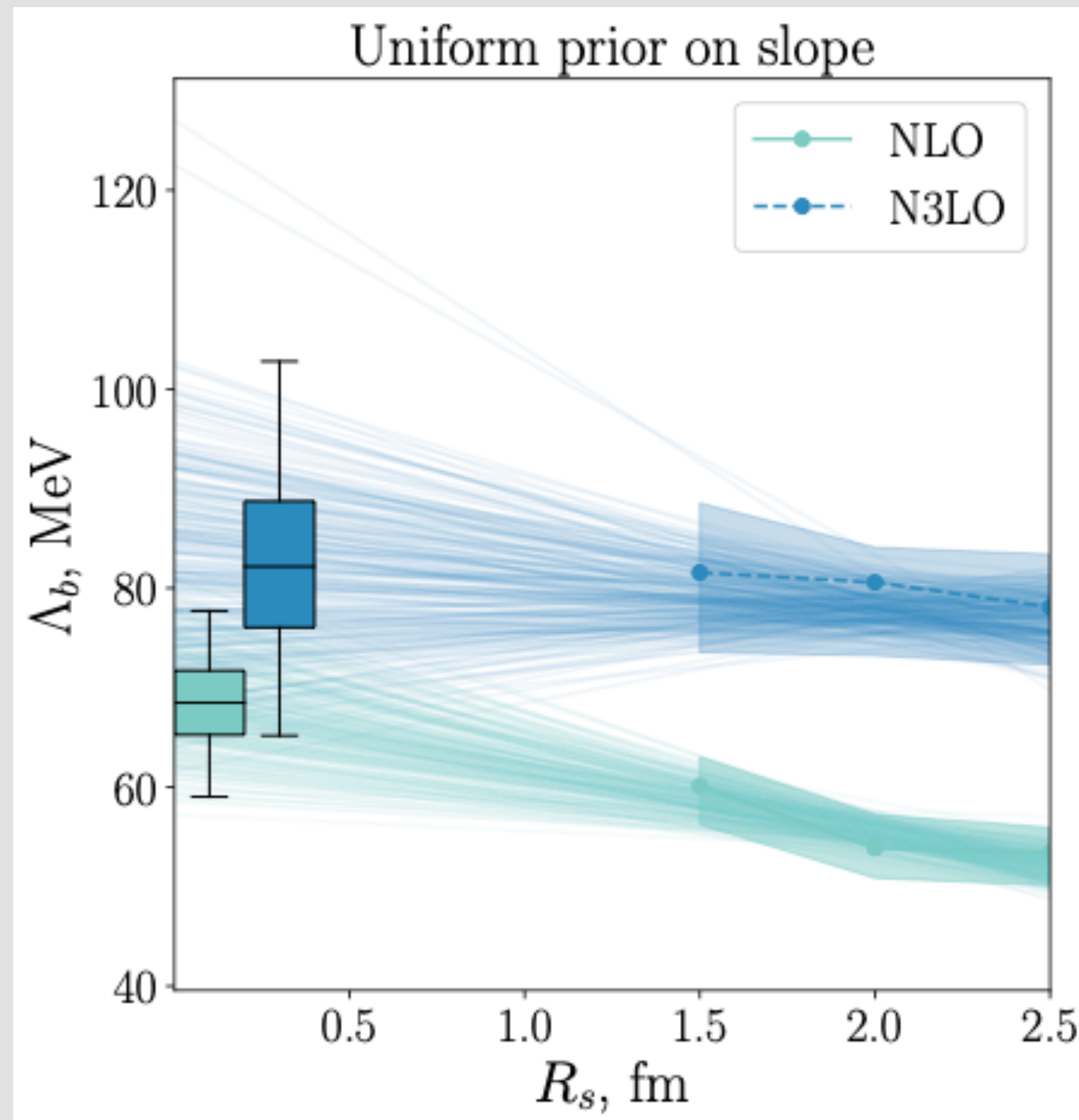


Extrapolation to Remove Artifacts



Why is there dependence on the order?

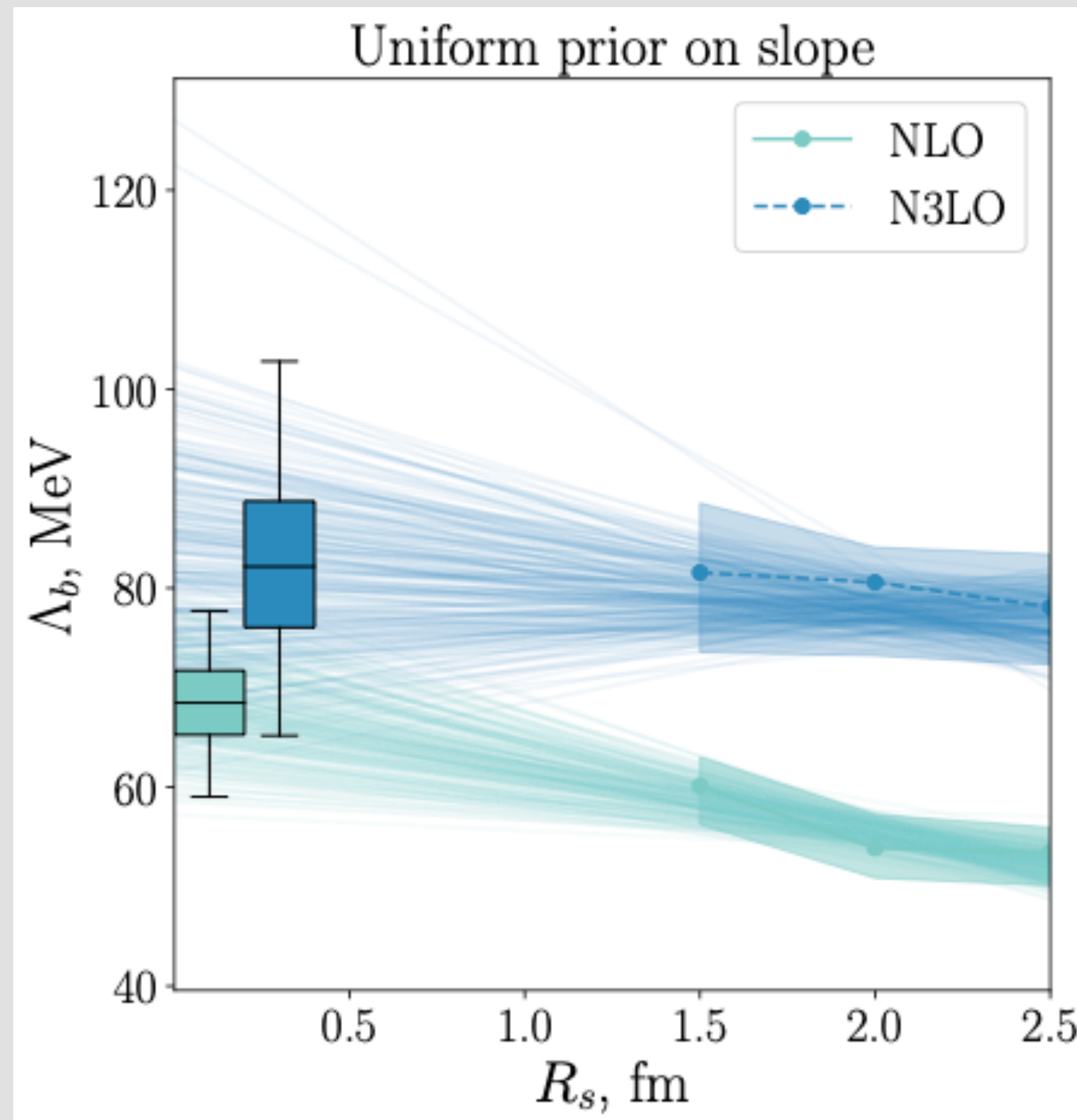
Extrapolation to Remove Artifacts



Why is there dependence on the order?

- Power counting?

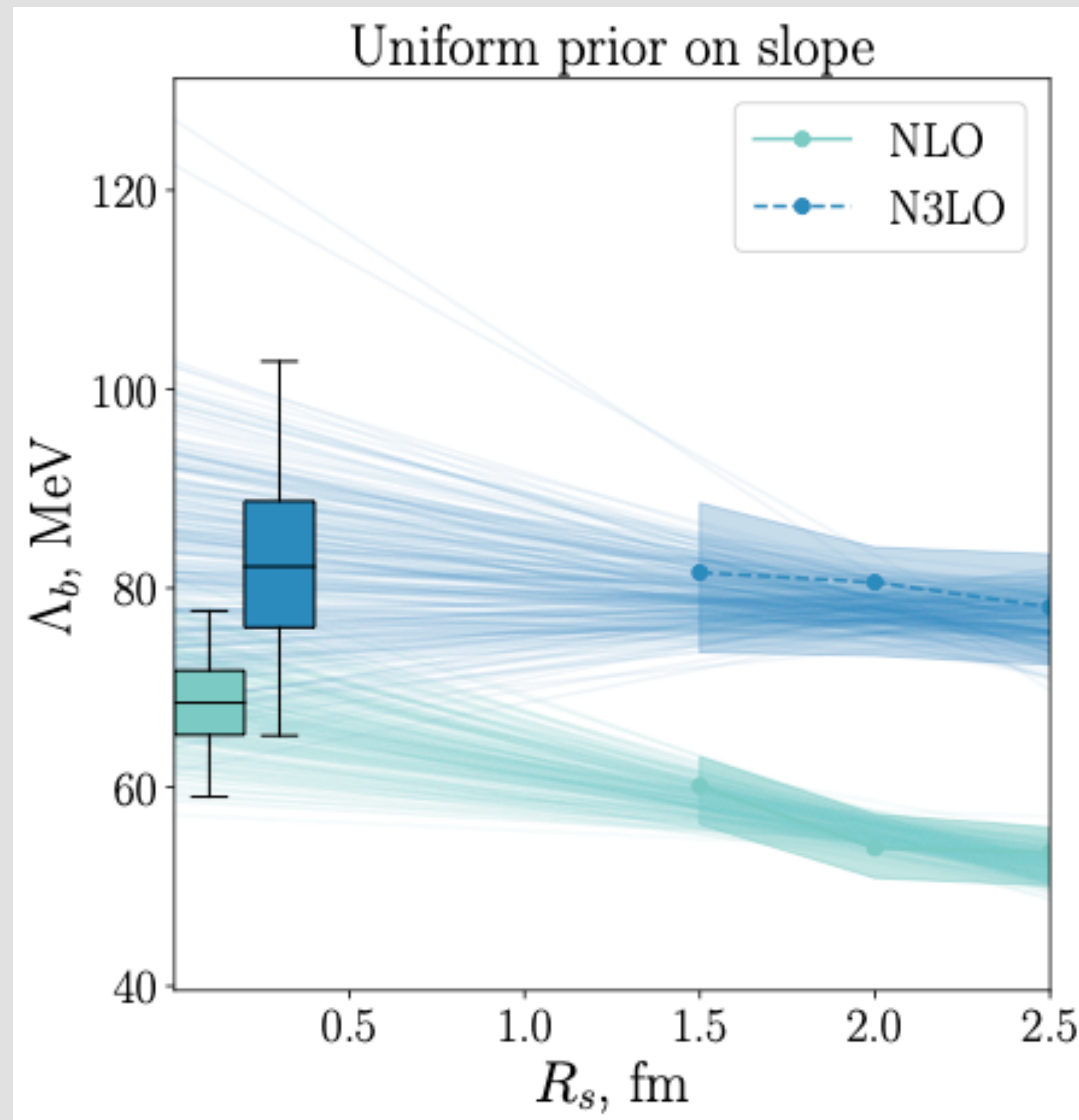
Extrapolation to Remove Artifacts



Why is there dependence on the order?

- Power counting?
- Flawed assumption of geometric series?

Extrapolation to Remove Artifacts



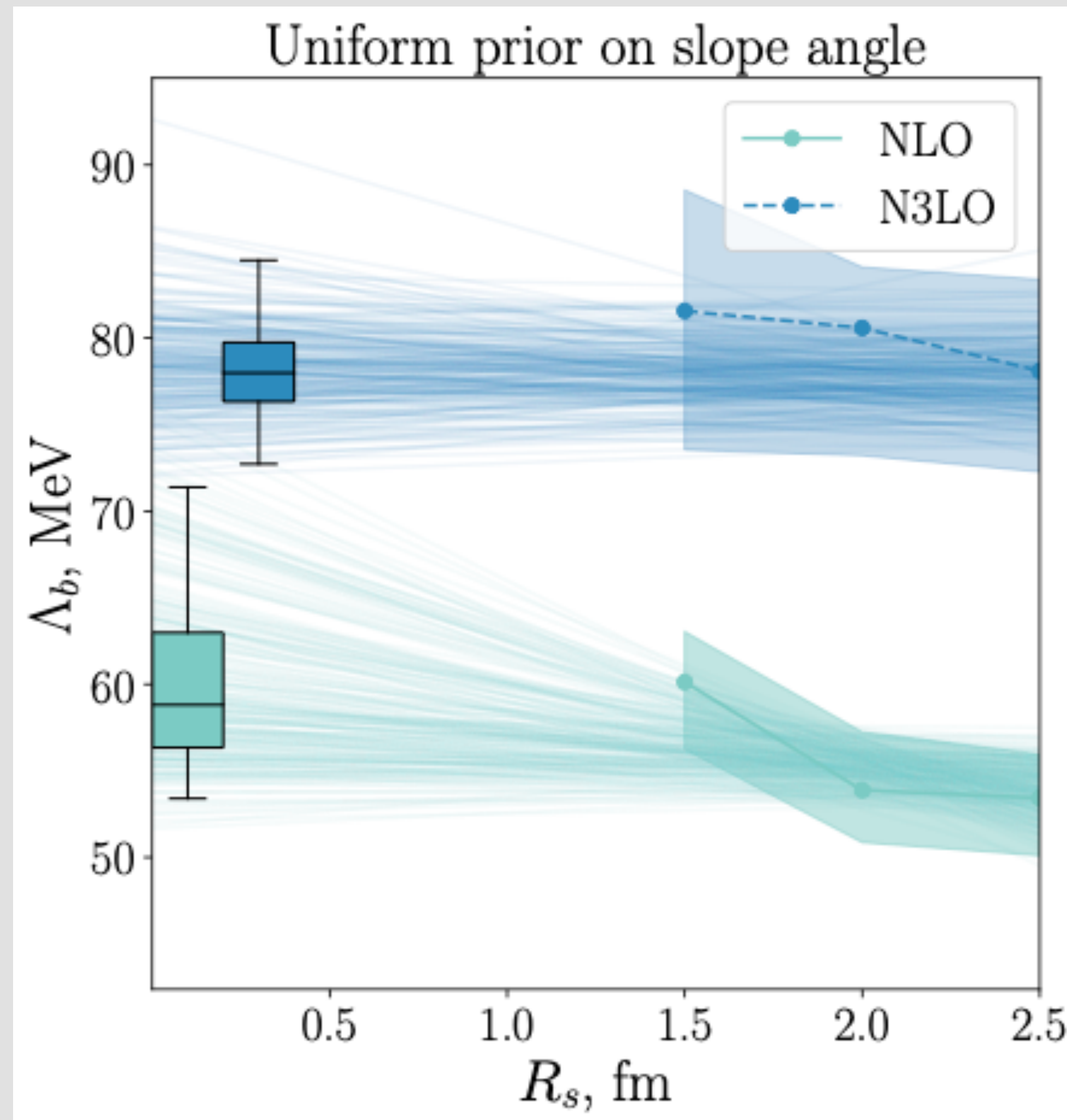
Why is there dependence on the order?

- Power counting?
- Flawed assumption of geometric series?
- Fierz transformation breaking?



Extrapolation to Remove Artifacts

Proper choice of prior!!!



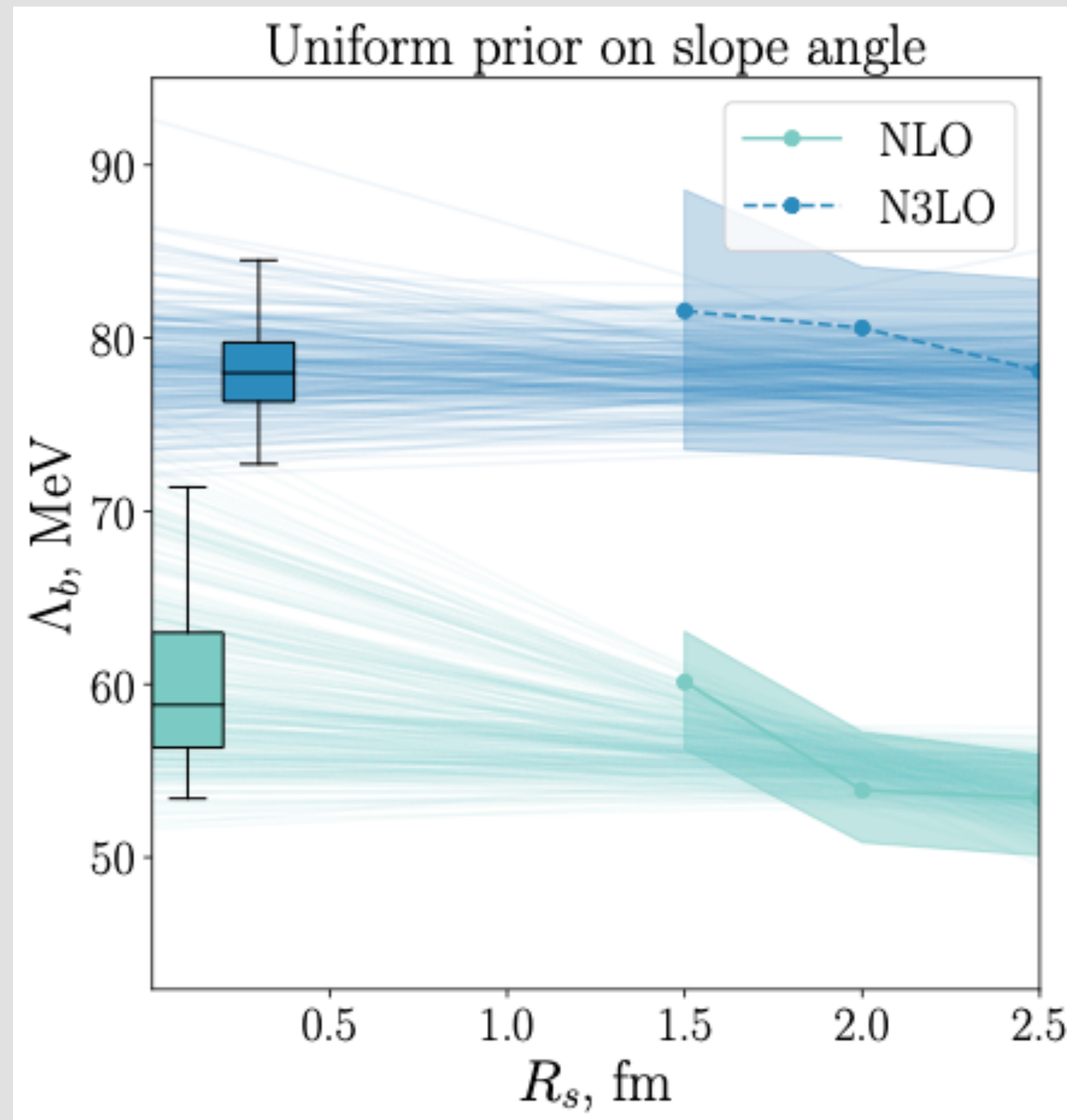
Why is there dependence on the order?

- Power counting?
- Flawed assumption of geometric series?
- ~~Fierz transformation breaking?~~



Extrapolation to Remove Artifacts

Proper choice of prior!!!



Why is there dependence on the order?

- Power counting?
- Flawed assumption of geometric series?
- ~~Fierz transformation breaking?~~

Unconstrained p -waves



Unconstrained p -waves



With $\Lambda_b \sim 50$ MeV, the max lab energy is given by

Unconstrained p -waves



With $\Lambda_b \sim 50$ MeV, the max lab energy is given by

$$\frac{p_{\text{c.m.}}^{(max)}}{\Lambda_b} = \frac{\sqrt{E_{\text{lab}}^{(max)} \mu}}{\Lambda_b} = 1 \Rightarrow E_{\text{lab}}^{(max)} = \frac{\Lambda_b^2}{\mu} \sim 5 \text{ MeV}$$



Unconstrained p -waves

With $\Lambda_b \sim 50$ MeV, the max lab energy is given by

$$\frac{p_{\text{c.m.}}^{(max)}}{\Lambda_b} = \frac{\sqrt{E_{\text{lab}}^{(max)} \mu}}{\Lambda_b} = 1 \Rightarrow E_{\text{lab}}^{(max)} = \frac{\Lambda_b^2}{\mu} \sim 5 \text{ MeV}$$

The Granada database has 4 data (polarized cross sections) up to 5 MeV that constrains 1P_1 and 3P_0 channels which have poor experimental and theoretical constraints.



Unconstrained p -waves

With $\Lambda_b \sim 50$ MeV, the max lab energy is given by

$$\frac{p_{\text{c.m.}}^{(max)}}{\Lambda_b} = \frac{\sqrt{E_{\text{lab}}^{(max)} \mu}}{\Lambda_b} = 1 \Rightarrow E_{\text{lab}}^{(max)} = \frac{\Lambda_b^2}{\mu} \sim 5 \text{ MeV}$$

The Granada database has 4 data (polarized cross sections) up to 5 MeV that constrains 1P_1 and 3P_0 channels which have poor experimental and theoretical constraints.

The data is telling us that our “Weinberg-ized” pionless EFT is predominantly s -wave physics!

Posterior Predictive Density



Posterior Predictive Density



We can now easily and **rigorously** propagate uncertainty to observable calculations.

Posterior Predictive Density



We can now easily and **rigorously** propagate uncertainty to observable calculations.

We calculate a **posterior predictive distribution (p.p.d.)** for the observables



Posterior Predictive Density

We can now easily and **rigorously** propagate uncertainty to observable calculations.

We calculate a **posterior predictive distribution (p.p.d.)** for the observables

$$\text{pr}(\vec{y}_{\text{th}} | \vec{y}, \vec{x}, I) = \int d\vec{a} d\bar{c}^2 d\Lambda_b \mathcal{N}(\vec{y}_{\text{th}}, \Sigma_{\text{th}}) \text{pr}(\vec{a}, \bar{c}^2, \Lambda_b | \vec{y}_{\text{exp}}, I)$$



Posterior Predictive Density

We can now easily and **rigorously** propagate uncertainty to observable calculations.

We calculate a **posterior predictive distribution (p.p.d.)** for the observables

$$\text{pr}(\vec{y}_{\text{th}} | \vec{y}, \vec{x}, I) = \int d\vec{a} d\bar{c}^2 d\Lambda_b \mathcal{N}(\vec{y}_{\text{th}}, \Sigma_{\text{th}}) \text{pr}(\vec{a}, \bar{c}^2, \Lambda_b | \vec{y}_{\text{exp}}, I)$$

which is done via sampling the posterior.



Posterior Predictive Density

We can now easily and **rigorously** propagate uncertainty to observable calculations.

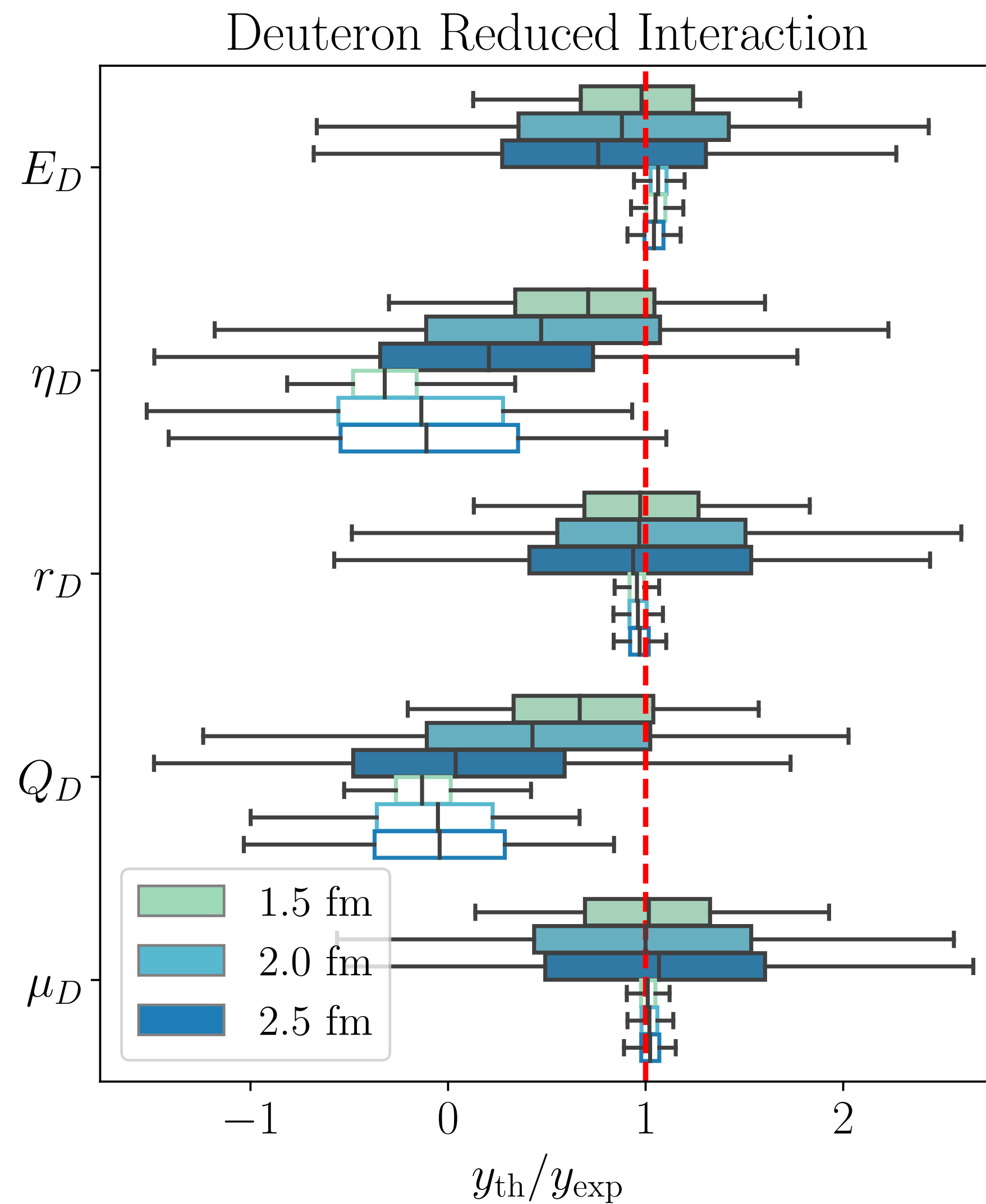
We calculate a **posterior predictive distribution (p.p.d.)** for the observables

$$\text{pr}(\vec{y}_{\text{th}} | \vec{y}, \vec{x}, I) = \int d\vec{a} d\bar{c}^2 d\Lambda_b \mathcal{N}(\vec{y}_{\text{th}}, \Sigma_{\text{th}}) \text{pr}(\vec{a}, \bar{c}^2, \Lambda_b | \vec{y}_{\text{exp}}, I)$$

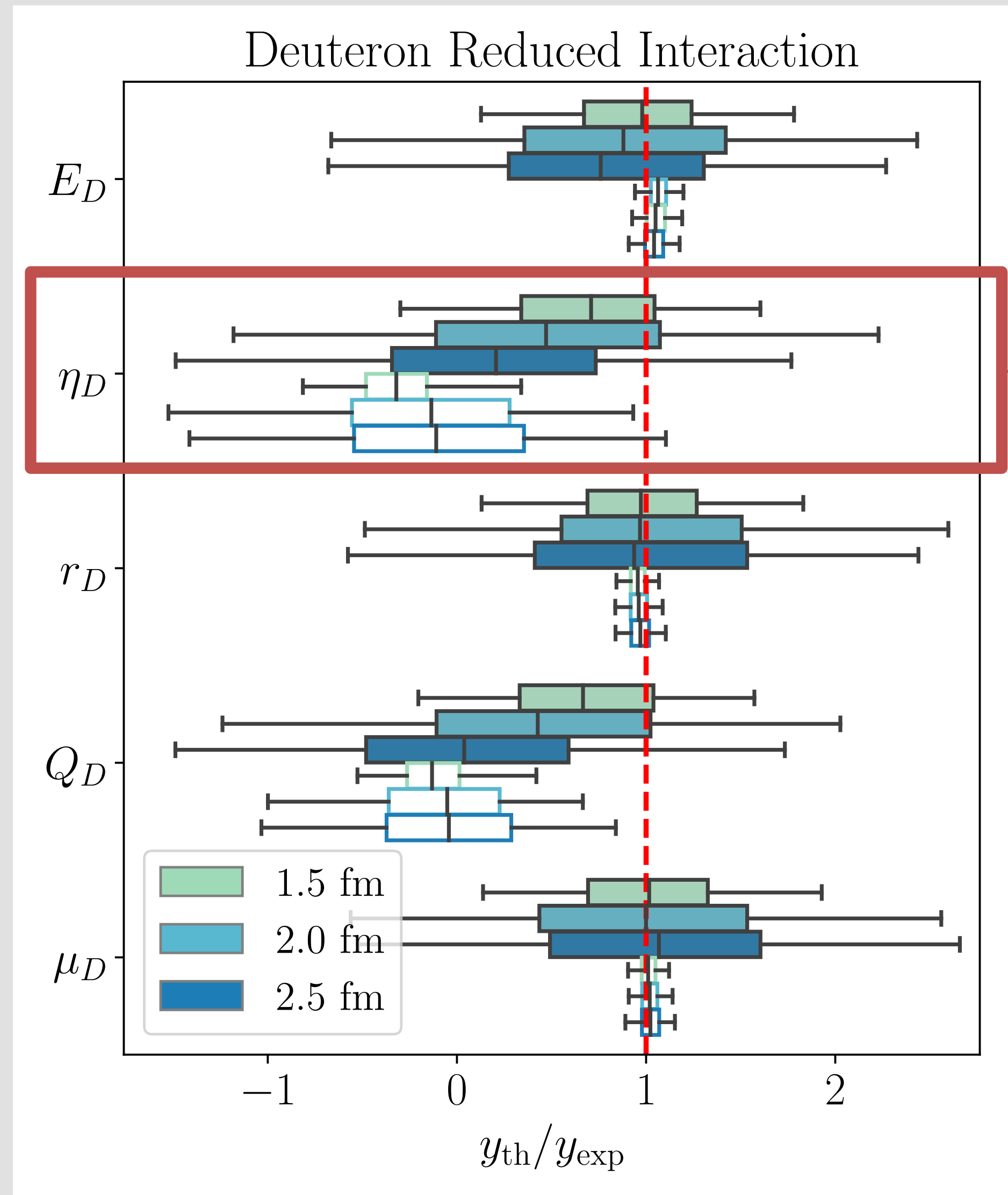
which is done via sampling the posterior.

This can be done for **any** calculation of nuclear observables.

Propagation of Errors for Deuteron

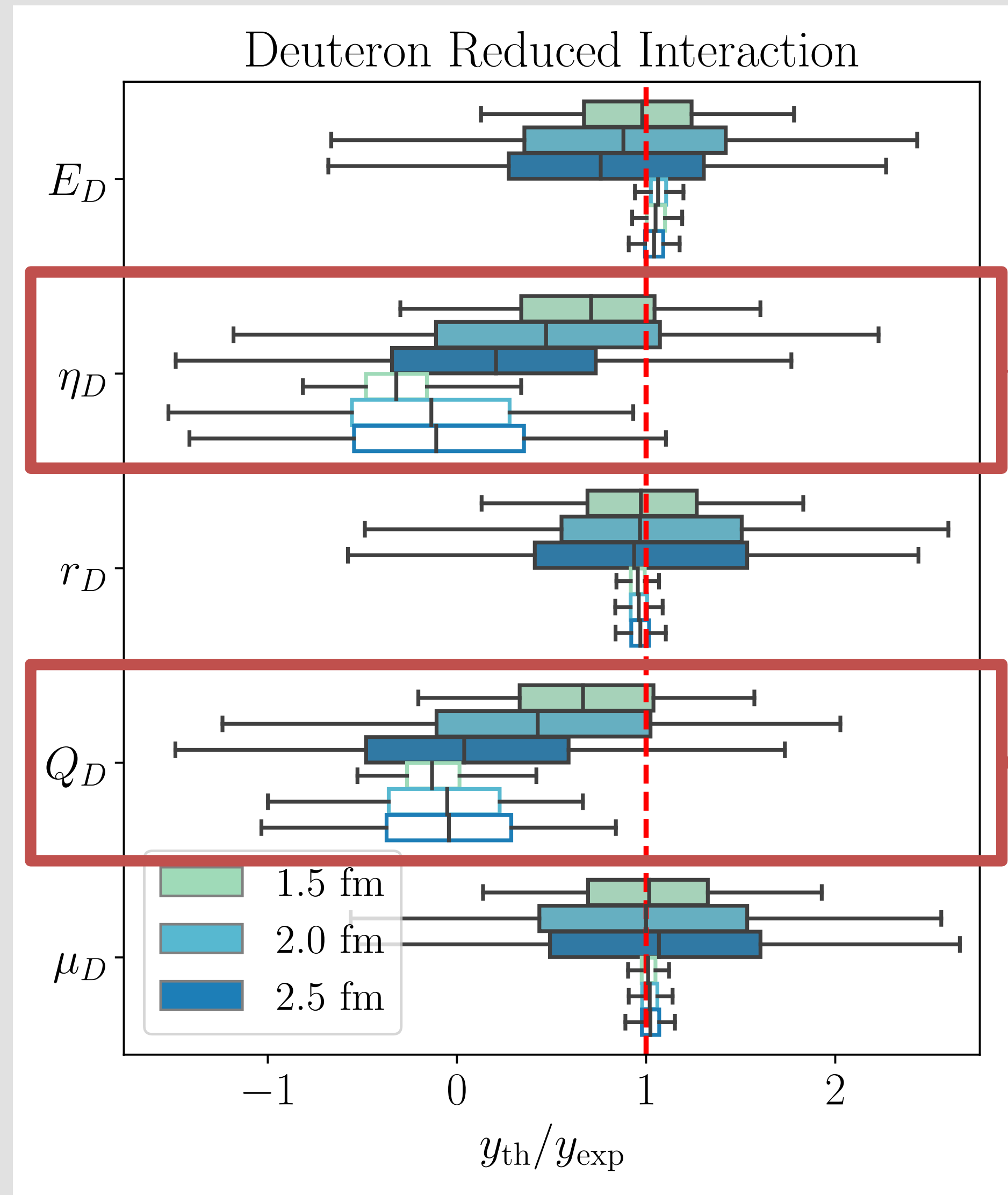


Propagation of Errors for Deuteron



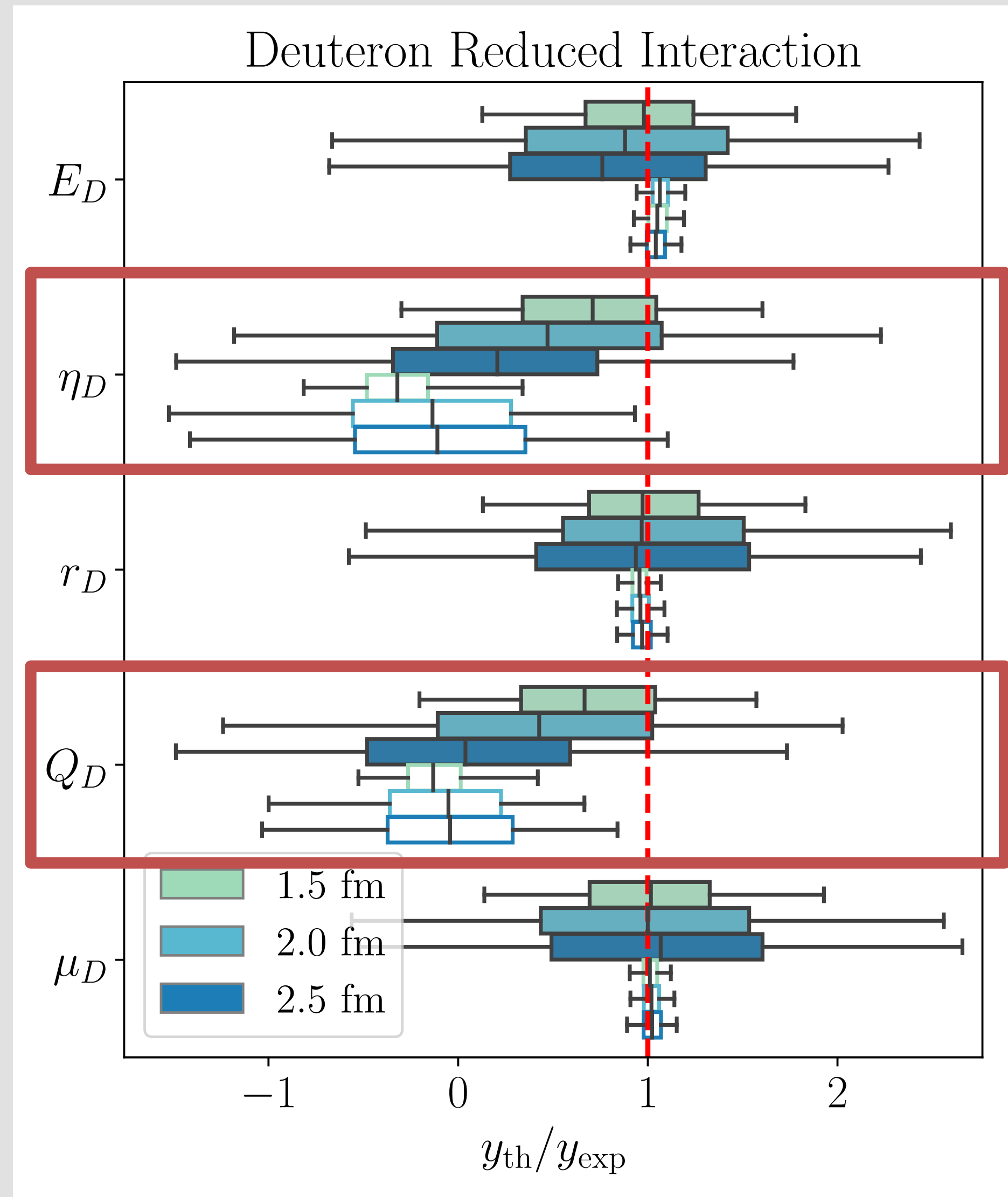
• Poorly constrained *d*-waves

Propagation of Errors for Deuteron



- Poorly constrained *d*-waves
- 2b corrections at $O(Q^5)$

Propagation of Errors for Deuteron



• Poorly constrained *d*-waves

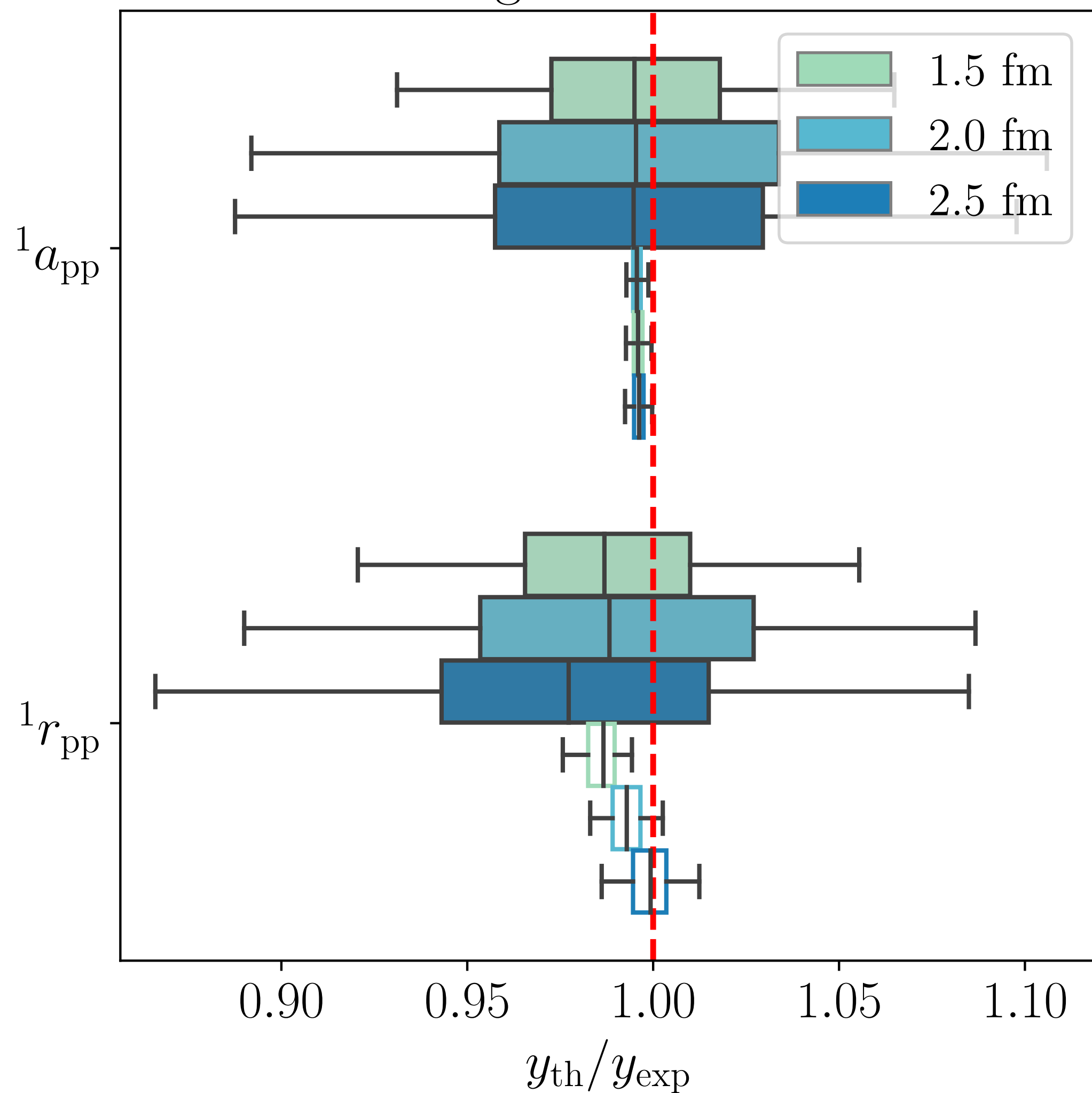
• 2b corrections at $O(Q^5)$

→ SHOULD BE
CONSISTENT WITH 0

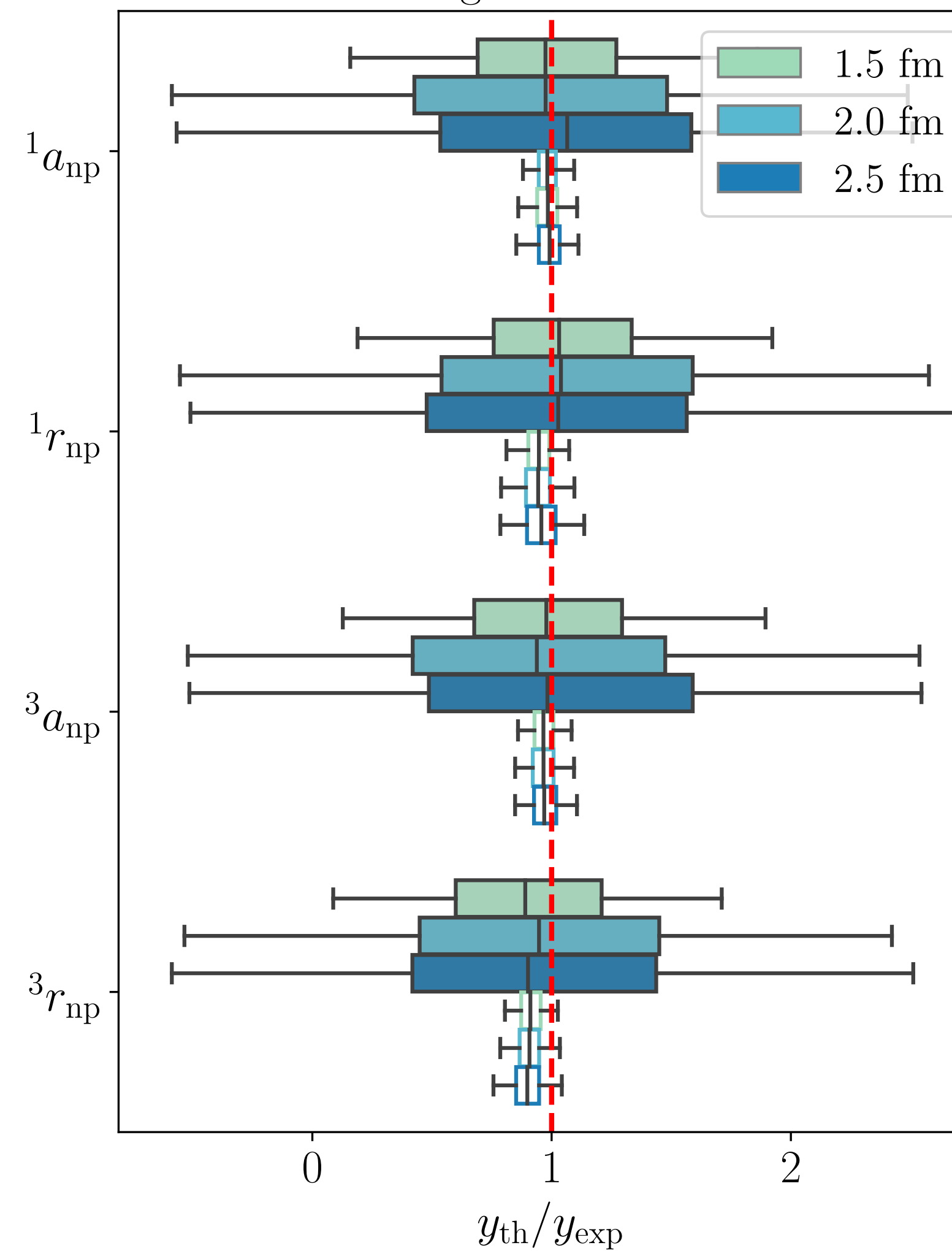
Propagation of Errors for ERPs



Effective Range Reduced Interaction



Effective Range Reduced Interaction



More Degrees of Freedom



More Degrees of Freedom



- For pion- and Δ -full interactions, we must look at higher energy data (~ 200 MeV)

More Degrees of Freedom



- For pion- and Δ -full interactions, we must look at higher energy data (~ 200 MeV)
- Emulation for calculation of scattering observables

More Degrees of Freedom



- For pion- and Δ -full interactions, we must look at higher energy data (~ 200 MeV)
- Emulation for calculation of scattering observables

$\sim 10^5$ parallel sample steps at ~ 4 min. per step

\rightarrow 280 days of wall time on an HPC

More Degrees of Freedom



- For pion- and Δ -full interactions, we must look at higher energy data (~ 200 MeV)
- Emulation for calculation of scattering observables



Ozge Surer
Miami University



Stefan Wild
LBNL



Matt Plumlee
Amazon



Pablo Giuliani
MSU/FRIB



Daniel Odell
SRNL

More Degrees of Freedom



- For pion- and Δ -full interactions, we must look at higher energy data (~ 200 MeV)
- Emulation for calculation of scattering observables



Ozge Surer

Miami University



Stefan Wild

LBNL



Matt Plumlee

Amazon



Pablo Giuliani

MSU/FRIB



Daniel Odell

SRNL

Gaussian Process Emulation

More Degrees of Freedom



- For pion- and Δ -full interactions, we must look at higher energy data (~ 200 MeV)
- Emulation for calculation of scattering observables



Ozge Surer

Miami University



Stefan Wild

LBNL



Matt Plumlee

Amazon



Pablo Giuliani

MSU/FRIB



Daniel Odell

SRNL

Gaussian Process Emulation

Reduced Basis Methods via
Galerkin Projection

Outline

- Motivation/Bayesian EFT Model Calibration
- BUQEYE Formalism
- Interaction Choice
- Results
- **Questions**

Open Questions



Open Questions



- Does proper inclusion of theoretical uncertainties in the calibration of EFT-inspired potentials bring all such models in agreement?

Open Questions



- Does proper inclusion of theoretical uncertainties in the calibration of EFT-inspired potentials bring all such models in agreement?
- Can Bayesian machinery help us identify consistent power counting? E.x. breakdown discrepancies across orders?

Open Questions

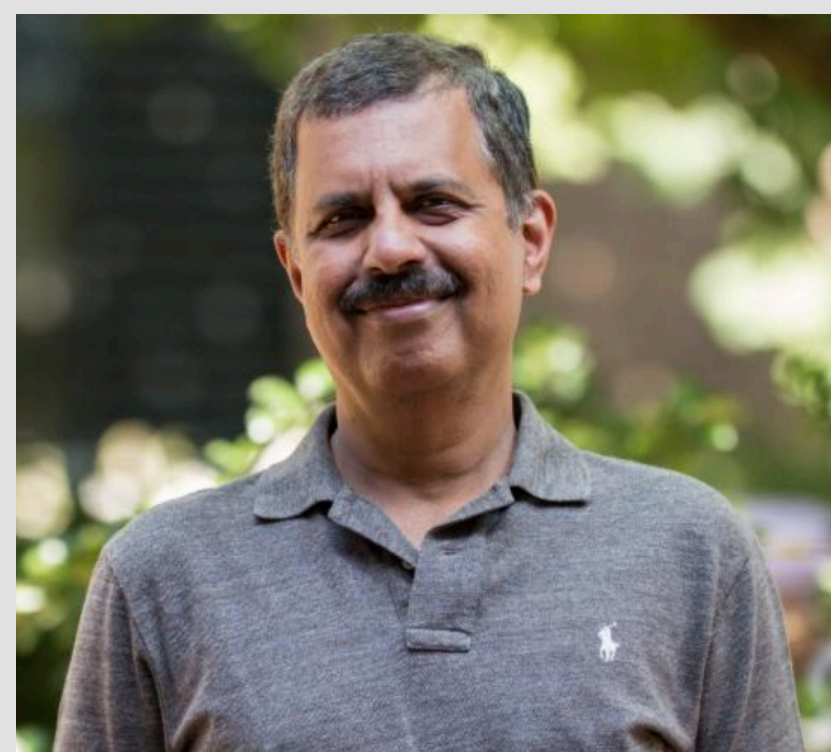


- Does proper inclusion of theoretical uncertainties in the calibration of EFT-inspired potentials bring all such models in agreement?
- Can Bayesian machinery help us identify consistent power counting? E.x. breakdown discrepancies across orders?
- Data can inform us about EFTs if we work consistently. How do we move towards consistent order-by-order inclusion of many-body forces and currents?

Acknowledgements

QMC@WashU

Piarulli (PI), Pastore (PI),
Novario (SS), Weiss(PD*),
Flores (PD), Chambers-Wall (GS)

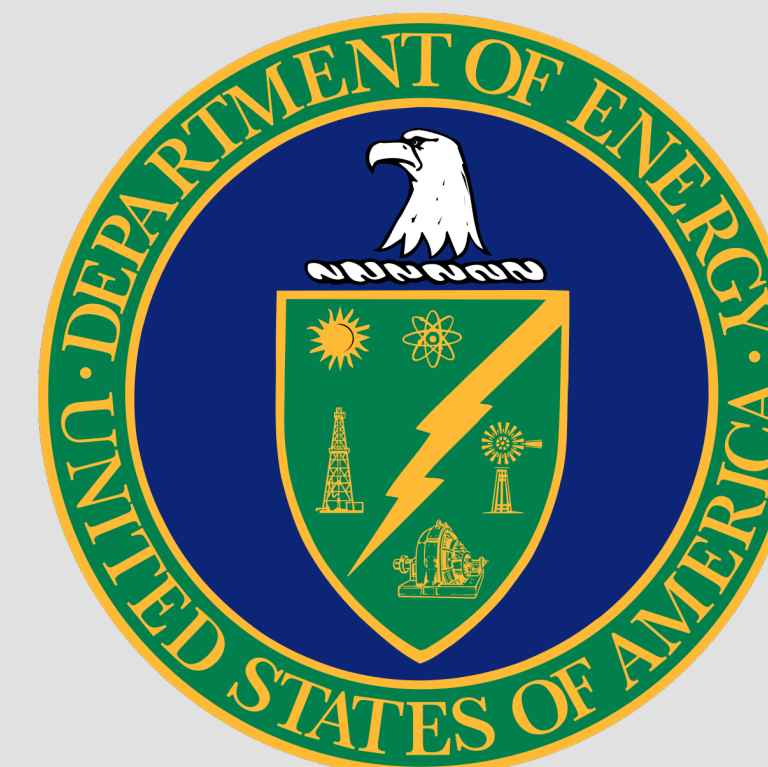


Sai Iyer
WashU

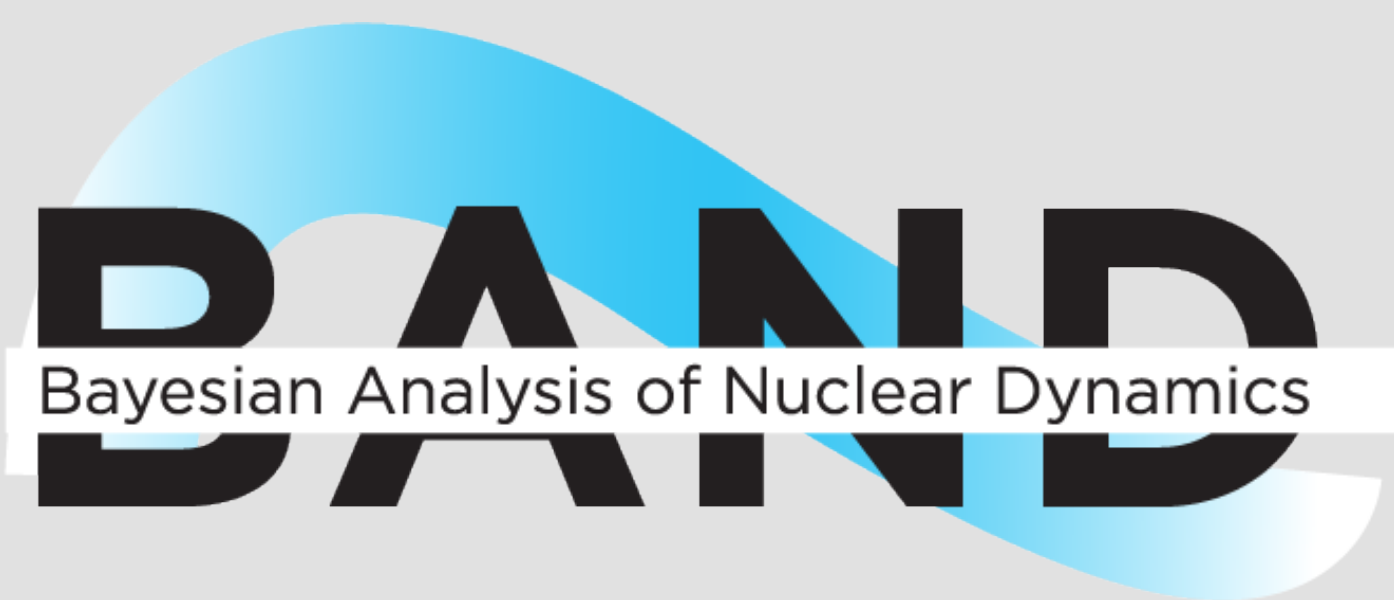
Computational Resources



Funding



Fellowship/Travel



Collaborations

NTNP

NUCLEI
Nuclear Computational Low-Energy Initiative
A SciDAC-5 Project